Generalized Linear Model Analysis

in Ecology

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Over the past two decades, advances in statistical analysis provided by generalized linear models (GzLMs) have improved the ability of researchers to investigate ecological processes. Relative to traditional general linear models (GLM), GzLMs provide greater flexibility due to their ability to analyze data with non-normal distributions. Considering the limitations of the GLM, we review the specific rationale of applying the GzLM in several fields of ecological research, and the number papers applying GzLMs. Through this review we revealed that GzLM has great potential for application across many fields of ecological research, and that there are an increasing number of research papers in ecology presenting results from GzLMs. We also analyze a series of exemplary ecological data sets using the GLM and GzLM, and used p-values to compare the relative sensitivity of the models. Finally, we assess the benefits and difficulties of applying GzLMs to ecological data. Because GzLMs allow for the specification of error structure, we found that GzLMs provided a much more appropriate fit to data sets with non-normal distributions, resulting in lower and more reliable p-values. Since most ecological data is non-normal, GzLMs provide a more effective analytical method than traditional linear models.

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Introduction

Data and statistical models of data are used in empirical sciences in order to gain a better understanding of processes and parameters. Statistical models provide a mathematical basis for the interpretation and examination of parameters and determine the roles and relative importance of different variables on a particular process¹. Statistical scientists have worked with ecologists for many years to improve the methods used to investigate, and thus better understand, ecological processes. Consequently, there are many modeling techniques available to ecological researchers. An important statistical development from the last 30 years is the introduction of the Generalized Linear Model $(GzLM)^2$, and the advancement and application of analysis provided through the GzLM regime in ecological research³. GzLMs are mathematical extensions of General Linear Models (GLM). GLMs provide familiar linear modeling and analysis of variance (ANOVA) tests which rely on traditional estimation techniques such as the least squares algorithm. The GzLM uses more flexible maximum likelihood parameter estimates; these estimates rely on an algorithm that iteratively uses a weighted version of least squares⁴. GzLMs are based on an assumed relationship, called a link function, between a linear predictor function of the explanatory variables and the mean of the response variable. Data are assumed to fall within one of several families of probability distributions, including normal, binomial, Poisson, negative binomial, or gamma². In comparison, GLMs are restricted to a normal error distribution and an identity link. Although a powerful approach when appropriately applied, GLMs are limited by the assumptions that errors are identically, independently, and normally distributed. Thus, the main improvements of GzLMs over GLMs are their ability to handle a larger class of error distributions, and

provide an efficient way of ensuring linearity and constraining the predictions to be within a range of possible values through the link function³. Thus, GzLMs provide a unified theory of modeling that encompasses the most important models for dealing with non-normal error structures. Since many of the data collected in ecological studies are poorly represented by normal distributions, GzLMs provide a flexible, suitable means for analyzing ecological relationships. Consequently, this approach has been extensively applied in many fields of ecological research, as evidenced by the growing number of published papers incorporating GzLMs.

Here, we evaluate the rationale, capabilities, and benefits of applying GzLMs within several fields of ecological research. We first present a literature review of the specific rationale for the use of GzLMs in various fields of ecological research, then proceed to evaluate the recent increase in usage of GzLMs across these fields. We then analyze several sample data sets from each field using both the GLM and GzLM, and evaluate the performance of each of these methods. We outline the advantages and disadvantages of these methods, and describe how the GzLM can be applied to optimize the analysis and maximize our understanding of ecological processes.

Field Rationale

The prevalence of GzLM in the existing literature is not evenly distributed across subdisciplines of ecology. Reasons for the common use or paucity of GzLM in a particular sub-discipline may be dependent on a variety of factors including types of data or lack of sophisticated analyses. This section addresses the potential for using GzLM and intrinsic/extrinsic limitations within six sub-disciplines of ecology:

Avian population monitoring

Monitoring long-term population change in birds is an integral part of effective conservation-oriented research and management. Since censuses of whole populations are often logistically impossible, population monitoring almost always relies on counts of subsets of a population. However, count data is often highly variable and overdispersed as a result of varying effort, missing data, observer differences, and actual natural variation. A number of analytical methods have been developed which attempt to deal with these complications, however, there is no consensus on which method is most suitable. The GzLM approach offers some promising solutions to these problems.

Boreal treeline dynamics

Many boreal researchers are studying the migration of the treeline, or tree invasion into alpine and tundra habitats. The GzLM, specifically logistic regression, is appropriate for this type of study. Logistic regression is useful for testing presence/absence, spatial distribution, and dominance of a species group as a function of biological and physical variables⁵⁻⁷. In addition, presence/absence data is commonly encountered in the field of treeline ecology, due to the advance of remote sensing and aerial photogrammetric technology used for monitoring these climatically sensitive areas⁸.

Marine bacteria

Bacteria are highly abundant in oceans, and function as a biological pump in mediating climate-active gases between ocean and atmosphere^{9, 10}. Studies dealing with aquatic bacterial abundance are essential to evaluate bacterial roles in biogeographic processes. Therefore, plenty of research has been done from lab-scaled experiments to global-scaled surveys of aquatic bacterial abundance. The overall understanding of these processes would be greatly increases with more sophisticated analysis. Most studies in this field use basic statistics, such as t-tests and Chi-square tests, to determine statistically significant effects. Model based statistics were seldom applied in this field. However, the more interesting question is which factors control the bacterial abundance and to which extent. Therefore, model based statistics will be much more useful, because it will report biological interest parameters, in addition to p-values.

The choice of GLM and GzLM to analyze bacterial abundance data cannot be determined without first knowing the distribution of the data. Bacterial pathogens shouldn't present in environment ubiquitously. Therefore, data collected about pathogen present and absent could be binomial distribution, which should be analyzed by GzLM.

Conservation biology involving vegetation

A primary focus in conservation biology, both floral and faunal, aims to determine the number of species in a given area. Whether the species of concern is a plant or whether the species distribution is conditional on a certain vegetation community, the use of GLM and GzLM are often used. It is therefore common that data of this nature is presence/absence or count data – both of which are most appropriately analyzed with

GzLM with an appropriate error structure. Estimates of survivorship, though underrepresented in the conservation literature, are of particular interest for population viability analyses and are arguably most appropriated analyzed using GzLM.

Marine and freshwater fish populations

Both marine and freshwater fish population studies often attempt to link species presence or abundance with specific environmental variables, such as water depth, temperature, or stream order. However, analyses of such data can often be difficult, due to the nature of sampling or the distribution of the fish population in question.

Marine studies rely heavily on trawl surveys to determine population distributions, which often produce a high number of zeroes in the dataset^{11, 12} due to the nature of schooling or aggregated fish^{13, 14}. Freshwater fish population studies typically use mark/recapture or electrofishing techniques, which can also give biased results. Fish captured in mark/recapture studies often show differing capture probabilities with time since being marked¹⁵ and electrofishing can produce significantly different counts of fish based on the methods used, such as the number of passes made¹⁶. These problems make analyzing data using the GLM difficult, as datasets are often highly skewed and nonnormal; by using the GzLM, these issues can typically be resolved¹⁷. Using binomial distributions, researchers can predict the probability of occurrence of fish in relation to environmental factors through the use of presence/absence response variables^{12, 14, 18-35}. The binomial approach can often be the better way to look at fish distribution, since it gives probability of capture^{15, 36}. This provides an idea of not just the amount of fish, but the magnitude or proportion of fish. The GzLM can also be used to estimate abundance from count data by using Poisson ^{37, 38} or gamma³⁹ error distributions, which account for the non-normal structure typical of fish count data. In the incidence of overdispersersion, a negative binomial distribution can be used to explain abundance^{13, 37, 40}. Some studies also use an integrated approach with GzLM's, by modeling probability of presence and abundance if present^{20, 41, 42}.

Field specific literature reviews for GzLM usage

Ecology is a broad field that encompasses many sub-disciplines of research, ranging from the study of microbes to entire ecosystems. In order to investigate the frequency of usage and specific applications of the GzLM in ecological studies, we therefore felt it was necessary to do literature reviews which were limited to several specific fields of ecology (conservational ecology, freshwater and marine fish population dynamics, avian population monitoring, boreal treeline expansion, and aquatic bacterial abundance). This method provided results that demonstrate the flexibility provided by the error structure in the GzLM and its ability to evaluate various types of biological data, along with the increasing frequency of usage of GzLMs in recent years.

All literature searches were carried out in the databases Web of Science and Biological Abstracts, using a pre-determined set of statistical query terms, along with ecological terms specific to each field (Table 1). It was possible that a greater number of references could have been found using further refined queries, but for the purposes of comparability across fields we chose to only use the pre-determined set of search terms. Search results were reviewed to determine if the GzLM was implemented, and if so, what type of data was being analyzed and what error structure was used (binomial, negative binomial, Poisson, Gamma). Articles in which the error structure was not stated were recorded as "Other" or on the basis of the extension used (Table 2). The total number of articles found per year was used to evaluate the frequency of GzLM usage over the past two decades (Fig. 1).

Of the six fields searched, the only one which did not produce any results was aquatic bacterial abundance. Within the articles found for conservational ecology, freshwater and marine fish population dynamics, and boreal treeline expansion, the most common error structure used was binomial, as presence/absence response variables were frequently analyzed; within these four fields, there was also some usage of the Poisson error structure, where response variables were count data (Table 2). The literature search within avian population trends, however, did not find any articles using binomial error – the majority of results in this field used Poisson (Table 2), as bird population trends are typically determined based on count data. Articles were also found which used negative binomial and Gamma error structures, but these were not as prevalent as binomial and Poisson (Table 2).

The literature reviews showed that GzLM usage has increased over the past two decades in all six fields searched (Fig. 1). The general trend produced when all results were totaled showed that GzLM usage was relatively low in the early 1990's, but increased steadily from the mid 1990's onward (Fig.1,a). Though this literature review only examined six fields of ecology, this trend appeared to occur across the five which produced results (Fig.1,b). If further refined searches were conducted and a larger number of fields surveyed, a more comprehensive picture of GzLM usage in ecology

would be obtained. However, the literature review of conservational ecology was fairly broad and produced a high number of results (131) which included a range of types of studies (Table 2), possibly indicating the trend for GzLM use in ecology in general.

 Table 1. Query terms input into Web of Science and Biological Abstracts for GzLM

 literature searches

| Field specific query terms | | | | |
|--|---|--|--|--|
| Conservational ecology of vegetation | Conservation, Ecology, Vegetation | | | |
| Freshwater fish population dynamics | Stream, Fish, Salmonid, Abundance, Distribution, Population | | | |
| Marine fish population dynamics | Demersal/Marine Fish Habitat, Abundance, Distribution | | | |
| Avian population monitoring | Birds, Aves, Population, Trends, Monitoring, Plots, Counts | | | |
| Treeline expansion Boreal, Treeline, Expansion | | | | |
| Aquatic bacterial abundance Aquatic, Bacteria, Abundance | | | | |
| Statistical query terms | | | | |
| Generalized and Generalised Linear Model(s), Generalized and Generalised Additive Model(s), Generalized and Generalised Estimating Equation, Logistic Regression, Poisson, Binomial, Gamma | | | | |

| Field | Data Type | Error Structure ^a / Extension ^b | References |
|--|-----------------------|--|---|
| Conservational Ecology of Vegetation | Presence/ Absence | Binomial | 31, 43-88,89-119,23, 120-152 |
| | Count | Poisson | 43, 153-164 |
| | ?? | GAM | 154, 165-167 |
| | Mixed | Other | 168, 169 |
| Freshwater | Presence / Absence | Binomial | 15, 22-29, 33, 35, 36, 39, 41, 42, 76, 125, 170-176 |
| Fish Population Dynamics | Count | Poisson | 39, 41 |
| · | Count | GAM | 177 |
| | Presence / Absence | Binomial | 12, 14, 18, 20, 21, 37, 176, 178 |
| | | Gamma | 20 |
| Marine Fish | Count | Poisson | 37, 179 |
| | | Negative Binomial | 13, 40 |
| | | GAM | 19, 180 |
| | | Poisson | 181-191 |
| Avian Population | Count | Negative Binomial | 58 |
| Monitoring | | GAM | 192-194 |
| | | Other | 184 |
| Boreal Treeline | Presence / Absence | Binomial | 5-7, 195-199 |
| Expansion | Count | Poisson | 198 |
| | Count | Other | 200 |

Table 2. Summary of field-specific usage of the GzLM

^a normal, binomial, Poisson, negative binomial, or gamma distributions ^b GAM, GzAM (Generalized Additive Model), GEE, GzAM (Generalized Estimating Equation)

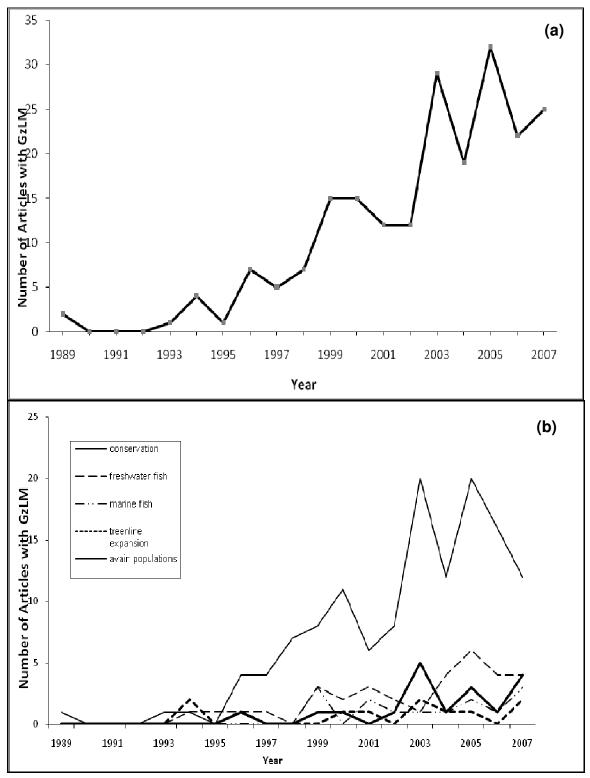


Figure 1. Frequency of GzLM usage resulting from literature searches of five ecological fields. **(a)** Frequency of usage in each field **(b)** Total incidence of usage, pooled from all searches.

Statistical analysis and summary of ecological data sets

GLMs and GzLMs were computed in the statistical programs SAS (9.1), R (2.6) and Minitab (13).

As seen in Table 3, one data set (B) could not be analyzed using the GLM, therefore this analysis is not included in the comparison of model results. From the remaining data sets, 13 of 21 parameters were significant when the GzLM was applied. Of these 13 parameters, 8 showed a decrease in p value from the GLM – 2 of which had not been significant until the application of the GzLM. The remaining 6 significant parameters had p values less than 0.001 and were not evaluated for change. Of the 8 nonsignificant parameters, 5 showed a decrease upon the application of the GzLM.

| | | GLM | | GzLM | |
|--|--|---|---|---|--|
| Study (Appendix) | Parameters | F(df) | Р | G(df) | р |
| | | | | error: Negativ | e binomial |
| Auklet Counts (A) | Gull | 36.9(1) | 2.00E-06 | 64.35(1) | 1.63E-08 |
| Aukiel Counts (A) | Species | 25.34(1) | 3.01E-05 | 32.43(1) | 1.60E-08 |
| | Gull*Species | 0.24(1) | 0.63 | 29.61(1) | 0.093 |
| | | | | error: Binomia | al |
| | Wetted Width | n/a | | 10.21(2) | 0.006 |
| | Depth | n/a | | 10.17(1) | 0.85 |
| Atlantic Salmon | % Rifle | n/a | | 6.19(1) | 0.046 |
| Distribution (B) | Width*Depth | n/a | | 4.33(2) | 0.40 |
| | Width*Rifle | n/a | | 2.34(2) | 0.37 |
| | Depth*Rifle | n/a | | 2.30(1) | 0.84 |
| Greenland Cod | | | | error: Negativ | |
| Counts (C) | Eelgrass | 3.37(2) | 0.037 | 7.54(2) | 0.023 |
| | | | | error: Binomia | |
| Juvenile Fish | Geology | 42.72(7) | <0.0001 | 319.43(7) | < 0.0001 |
| Distribution (D) | Time | 17.73(1) | < 0.0001 | 17.89(1) | |
| | | | | error: Negativ | |
| Seed Removal by | Species | 18.11(4) | 9.27E-11 | 53.14(4) | 1.77E-10 |
| Red Ant (E) | Mound | 2.17(1) | 0.14 | 0.74(1) | 0.39 |
| | liounu | | 0111 | error: Begativ | |
| | Fire | 1.22(1) | 0.29 | 201.01(1) | 0.069 |
| Red Oak | Basal Area | 3.19(1) | 0.093 | 136.80(1) | 0.0002 |
| Recruitment (F) | Down Woody | 5.15(1) | 0.055 | 190.00(1) | 0.0002 |
| | Debris | 0.49(1) | 0.50 | 123.01(1) | 0.089 |
| | | | | * * | |
| Seedling | | 126.33 | | | |
| herbivory (G) | groundcover | (2) | 0.008 | 14.06(2) | 0.0009 |
| | pesticide | 90.75 (1) | 0.011 | 5.29(1) | 0.022 |
| Cood availation | | | | error: Negativ | e binomial |
| | groundcover | 11.04 (2) | 0.083 | 27.80(2) | < 0.0001 |
| (П) | pesticide | 2.29 (1) | 0.27 | 1.63(1) | 0.20 |
| Common Murre | | | | error: Negativ | e binomial |
| plot counts (I) | Year | 23 (326) | 2.00E-06 | | |
| Thick-billed | | | | error: Negative binomial | |
| murre plot counts | | 284.56 | | | |
| (I) | Year | (676) | <2.2e-16 | 236.03 (1) | 2.88E-53 |
| De eterriel | | | | error: Gamma | |
| | water masses | 19.57(1) | 4.00E-05 | 13.69(1) | 2.09E-07 |
| abundance (J) | pН | 0.11(1) | 0.74 | 0.24(1) | 0.44 |
| | · | | | error: Poissor | |
| Colony counts (K) | dishes | 1.36(1) | 0.25 | 1.21(1) | 0.27 |
| | quadratic | 1.55(1) | 0.22 | 1.37(1) | 0.25 |
| herbivory (G) Seed predation (H) Common Murre plot counts (I) Thick-billed murre plot counts | groundcover pesticide groundcover pesticide Year | 126.33 (2) 90.75 (1) 11.04 (2) 2.29 (1) 23 (326) 284.56 | 0.008 0.011 0.083 0.27 2.00E-06 | error: Binomia 14.06(2) 5.29(1) error: Negativ 27.80(2) 1.63(1) error: Negativ 19.13 (1) error: Negativ 236.03 (1) | al 0.0009 0.022 e binomial <0.0001 0.20 e binomial 1.22E-05 e binomial 2.88E-53 |

 Table 3. Comparison of GLM and GzLM data analysis results (see

 Appendices A-K for detailed results)

Advantages and disadvantages to applying the GzLM

Over the duration of the Biology 7932 graduate course, we developed our understanding of the generalized linear model and how to apply it to data with a non-normal error distribution. From our analyses, we have discovered some general problems that arise when using the GzLM, some of them particular to our exemplary data sets. Here, we discuss advantages and disadvantages inherent to applying the GzLM to data with Poisson, negative binomial, binomial and gamma distributions. Some were problems directly encountered during the analysis of the exemplary data sets, and some are more general limitations drawn from various ecological studies.

Poisson and negative binomial distributions

The generalized linear model can be applied to data with both the Poisson and negative binomial distribution. The Poisson distribution is usually linked to count data, which commonly appears in ecological literature; count data are generated through studies that gather model information through mark-recapture experiments¹⁷⁰, site-specific captures such as trawls¹³ and plot counts, a common method of evaluating seabird populations^{172, 181}

Population monitoring of birds almost always relies on surveys of a sample population ²⁰¹. This is because it is not feasible to census the entire population of most birds. Consequently, censuses of large colonies are usually based on sub-samples of the population²⁰². In this section we review the analysis of long-term count data from census plots at Cape St. Mary's, Newfoundland, 1980-2006 (Appendix I).

The analysis of count data is often complicated due to the subjective nature of trend estimation and high inherent variation. There are many statistical techniques that may be employed to analyze count data, however, there is little consensus regarding the most suitable method ²⁰¹. Linear regression is one of the oldest statistical technique, and has been long been used in ecological research to test trends. When a linear regression model was fit to the Cape St. Mary's common murre plot data, we found that it preformed poorly, exhibiting non-normal, non-independent and heterogeneous residuals.

A second option is to run a GzLM with a Poisson distribution (Poisson regression). Poisson regression is a frequently used method for count data. A key feature of Poisson distribution is that it assumes that as the mean increases, the variance increases – which is a frequent characteristic of count data⁴. Nevertheless, it appears that the murre count observations exceeded the amount of variation predicted by Poisson, whereby the estimated overdispersion parameter (ϕ [Null deviance/df]) was much greater than 1 (ϕ = 8.9). The Poisson distribution may have fit the data poorly since it is designed to be used for counts of events that occur randomly over time or space²⁰³. The murre plot count data are neither temporally or spatially independent since counts are conducted at the same time for all plots and the same plots are monitored across years.

The negative binomial is a distribution related to Poisson, however it includes an extra parameter (dispersion parameter) which allows the variance to exceed the mean²⁰⁴. Counts of populations are often fitted well by the negative binomial distribution²⁰⁵. White and Bennetts²⁰⁶ suggest that ecological count data likely exhibit a negative binomial distribution more frequently than Poisson or normal distributions. When a

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GzLM with a negative binomial distribution was fitted to the murre plot data, the overdispersion parameter was much closer to 1 ($\phi = 1.1$).

As apparent from the analysis of the Cape St. Mary's murre plot trend data, one of the main advantages of GzLMs over GLMs is that they do not force data into unnatural scales by allowing for non-consistent variance structures in the data. Linear regression models are limited by the assumptions that the errors are identical, independently, and normally distributed. In the case of the murre count data, the GzLM model allowed for the selection of different distributions such as Poisson or negative binomial, which better suited the data. Within the options available in the GzLM, the negative binomial distribution appeared to best suit the data. The dispersion parameter accounted for the extra variance which exceeded the assumptions in the Poisson regression.

The main disadvantage of the Poisson distribution is that the scope of data that can be analyzed using this distribution is limited by the assumption that the mean increases with the variance. The Poisson distribution do not account for any extra variation in the data, thus in such cases, a negative binomial distribution may be more appropriate. The primary disadvantage of using negative binomial regression is that there are fewer programs that are capable of building a GzLM with a negative binomial distribution. Another disadvantage of the negative binomial distribution is the inclusion of the dispersion parameter. With an increase in parameters, there is a decrease in precision and one may run the risk of overfitting the data, thus less weight should be placed on models with a higher number of parameters¹³⁶. One general disadvantage of the GzLM is the lack of diagnostic plots available to assess the model. One specific problem relating to the count data is in the estimation and assessment of the overdispersion parameter – there appears to be no standard method of calculation, or a definition of how much the estimate needs to deviate from 1 for the data to be considered overdispersed.

Binomial distributions

Binomial data is common in ecological modelling. It often appears as presence / absence or risk data, which is binary in nature, so unlike other non-normal distributions, it is usually relatively easy to determine when to apply the GzLM with a binomial distribution. When the data are binary and the distribution is accepted as non-Gaussian, the data are usually analysed using logistic regression. Logistic regression is a GzLM analysis that is often used in ecological research; it has been applied for evaluating habitat²⁰⁷, assessing risk²⁰⁸, and predicting the distribution of species and vegetation groups^{6, 209}.

The seedling herbivory data collected in the Mealy Mountains during the summer of 2007 (See Appendix G for details) demonstrate a clear example of a binomial response variable (in this case, herbivory occurs or does not occur for each seedling). The purpose of the study was to determine whether the odds of herbivory were influenced by a variety of environmental factors; since the response variable is binary, we assumed a binomial distribution and applied a logistic regression model.

A limitation of the GzLM was encountered when analysing the binomial seedling herbivory data; here, we review the difficulty in performing a prospective power analysis for a logistic regression model. A prospective power analysis is used to help the researcher design their study, ensuring that the effect size, sample size, and level of precision are all sufficiently large to generate an experiment with sufficient statistical power²¹⁰. A prospective power analysis was performed using the SAS macro UnifyPow on the herbivory logistic regression model in order to determine the increase in sample size needed to maintain the power of the study through a second experimental season; since seedlings were being removed through herbivory, the sample size (and thus the power) was decreasing over time. However, according to the power analysis, the power of the study remained equal regardless of any increases in the initial population size. The problems of prospective power analyses and their application to binomial GzLMs are not well documented in the ecological literature. Power analyses for logistic regression models are generally used in medical studies²¹¹, but these cannot be easily compared to ecological studies.

In general, there are some limitations to the application of the GzLM for analysing binary data generated by ecological studies. In ecology and conservation biology, logistic regression models are relatively common, and often used for modelling spatial species distributions under different environmental conditions^{6, 199}. However, these models must assume that the species is in a state of pseudo-equilibrium with the environment; therefore, logistic regression models cannot effectively identify environmental factors responsible for distribution when a species is not in equilibrium, or still expanding its range³. Furthermore, GzLMs and GAMs are usually based on empirical data collected from a particular region, thus incorporating the biotic interactions and random effects that are characteristic to the area. This means that predictive GzLMs designed for one region usually cannot be applied on a wider scale; the predictive power over a broad spatial scale is usually low^{3, 212}. Another general disadvantage of logistic regression analysis is the tendency of the model to strongly underestimate the probability of rare events²¹³.

Gamma distributions

The gamma distribution also occurs in ecological studies, though much less frequently than binomial and Poisson according to our literature search. Thus, there is less information available about its application to ecological data. The gamma distribution is a 2-parameter frequency distribution given by the equation:

$$f(x) = \frac{1}{\beta^{\gamma} \Gamma(\gamma)} x^{\gamma - 1} e^{-x/\beta}; \beta > 0, \gamma > 0$$

Gamma distribution has a zero lower bound and is unlimited on the right. It is positively skewed, with the amount of skew depending inversely on the shape factor γ . The gamma distribution is closely related to the Chi-square distribution, for $\chi^2/2$ is a gamma variate²¹⁴.

In biological ecology, data distributions are highly variable; as we have observed, most data are not normally distributed. The general linear model is usually applied when the distribution is gamma; this occurs partially because basic linear statistics (based on the normal distribution) have been traditionally and widely taught. The major advantage of using the gamma distribution is that it simplifies the interpretation of the model; if you force the model to fit the normal distribution where gamma is applicable, the variables need to be transformed, which complicates the biological interpretation of the parameters. Gamma distribution helps expand the capacity of data analysis without using transformation. The major disadvantage of using the gamma distribution is common to many applications of the generalized linear model; in order to analyse a model using gamma distribution, you need a powerful statistical software package and also relatively advanced knowledge of statistics.

Conclusions

The GzLM has become increasingly common in ecological studies over the past two decades. This is largely due to the flexibility of the GzLM when compared to the GLM, in that it does not assume normal distributions. The nature of ecological data often produces data which is highly varied, has counts with many zeros, is over dispersed, skewed, or more suited to analysis of presence versus absence. The specification of the error structure in the GzLM allows ecological researchers to create models which can account for the nature of these datasets and provide more meaningful statistical analysis. Despite these advantages of GzLMs, their use has not become widespread until recently. It is likely that as the acceptance and discussion of the GzLM increases, its role as the most appropriate model for ecological data analysis will be recognized in the field.

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Appendix A

Auklet Count Data Christina Bourne

Research Question: Do least auklets (*Aethia pusilla*) and crested auklet (*Aethia cristatella*) surface counts increase with time since predator disturbance?

Data: Surface counts of auklets on a 10x10m plot, Buldir Island, Alaska. 4 hours of observations per day with counts made every 15 minutes, time of all predator disturbances recorded. Total of 25 days used in analysis; 375 counts per species. All data collected on Buldir Island, Alaska during the summer of 2004 by Christina Bourne.

Model 1: General Linear Model; log transformed mean surface counts

Formal Model: $C = \beta_0 + \beta_G * G + \beta_S * S + \beta_{G*S} * G*S + res$

Table A1: ANOVA

| | • · · · | | | |
|-----------|---------|----------------|---------|-----------|
| | Df | Sum Sq Mean Sq | F value | Pr(>F) |
| Gull | 1 | 4.2747 4.2747 | 36.9793 | 1.998e-06 |
| Sp | 1 | 2.9289 2.9289 | 25.3375 | 3.075e-05 |
| Gull:Sp | 1 | 0.0278 0.0278 | 0.2409 | 0.6277 |
| Residuals | 26 | 3.0055 0.1156 | | |

Model 2: Generalized Linear Model; error = Negative Binomial

| Table A2: GzLM summary | | | | | | |
|--------------------------------|----------|------------|---------|-----------|--|--|
| | Estimate | Std. Error | z value | Pr(> z) | | |
| (Intercept) | 3.92610 | 0.82073 | 4.784 | 1.72e-06 | | |
| Gull | 0.02816 | 0.08261 | 0.341 | 0.733176 | | |
| Sp | -2.23221 | 0.60250 | -3.705 | 0.000211 | | |
| Gull:Sp | 0.10009 | 0.05835 | 1.716 | 0.086249. | | |
| 1 | | | | | | |
| Table A3: GzLM Chi-square test | | | | | | |

| Table A5: GZLIVI CIII-square test | | | | | | |
|-----------------------------------|----------|---|---|--|--|--|
| Df | Deviance | Resid. | Df Resid. | Dev P(> Chil) | | |
| | | 29 | 96.240 | | | |
| 1 | 31.887 | 28 | 64.352 | 1.634e-08 | | |
| 1 | 31.927 | 27 | 32.425 | 1.600e-08 | | |
| 1 | 2.816 | 26 | 29.609 | 0.093 | | |
| | | Df Deviance 1 31.887 1 31.927 | Df Deviance Resid. 29 1 31.887 28 1 31.927 27 | Df Deviance Resid. Df Resid. 29 96.240 1 31.887 28 64.352 1 31.927 27 32.425 | | |

Advantages of GzLM:

- does not assume that model mean and variance are equal (for over dispersed count data they are not); uses additional parameter to adjust the variance independently of the mean (Hinz and Gurland, 1968).
- can use actual count data (as opposed to transformed done to stabilize the error variance) which potentially increases interpretive value of results

• provides lower p values for the same data set than general linear model, indicating greater statistical power

Disadvantages of GzLM

- can be a more difficult model to implement in some statistical programs
- has an extra parameter

Appendix B

Salmon Habitat Distribution Christina Bourne

Research Question: Is there a relationship between stream characteristics (width, depth, rifle) and Atlantic salmon (*Salmo salar*) distribution?

Data: 53 salmon counts from electrofishing of 12 streams in Terra Nova National Park, NL, in the summers of 2005 and 2006. Average wetted width, average depth and estimated percent rifle for each site was recorded.

* Data provided by Dr. Dave Cote, aquatic ecologist, Terra Nova National Park

Model 1: General Linear Model

Formal Model: C = $\beta_0 + \beta_w *W + \beta_D *D + \beta_R *R + \beta_{W*D}*W*D + \beta_{W*R}*W*R + \beta_{D*R}*D*R + res$

This model did not meet the assumptions for the general linear model when counts, mean counts and log transformed counts (Fig. B1) were used as response variables.

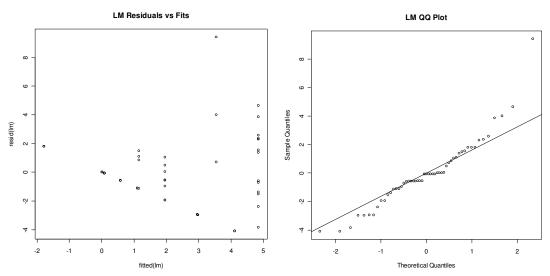


Fig B1. Residuals vs fits plot and qq plot for log transformed salmon counts

Model 2: Generalized Linear Model; error = binomial

Formal Model: C = $\beta_0 + \beta_w *W + \beta_D *D + \beta_R *R + \beta_{W*D}*W*D + \beta_{W*R}*W*R + \beta_{D*R}*D*R + res$

Table B2: GzLM summary

| Table D2: G2LM Summary | | | | | | |
|------------------------|----------|------------|------------------|--|--|--|
| | Estimate | Std. Error | z value Pr(> z) | | | |
| (Intercept) | -0.63011 | 1.12128 | -0.562 0.574 | | | |
| Width4m | -0.02326 | 1.46152 | -0.016 0.987 | | | |
| Width6m | .23668 | 1.42126 | 0.870 0.384 | | | |
| Depth20cm | -0.12742 | 1.45440 | -0.088 0.930 | | | |
| Rifle100% | 0.54515 | 1.51172 | 0.361 0.718 | | | |
| Width4m:Depth20cm | -1.36015 | 1.90017 | -0.716 0.474 | | | |
| Width6m:Depth20cm | 1.06101 | 1.70333 | 0.623 0.533 | | | |
| Width4m:Rifle100% | 2.07279 | 1.95559 | 1.060 0.289 | | | |
| Width6m:Rifle100% | -0.27718 | 1.70744 | -0.162 0.871 | | | |
| Depth20cm:Rifle100% | 0.29735 | 1.47303 | 0.202 0.840 | | | |
| | | | | | | |

| Table B3: GzLM Chi-square test | | | | | | |
|--------------------------------|----|----------|-----------|------------|-----------|--|
| | Df | Deviance | Resid. Df | Resid. Dev | P(> Chil) | |
| NULL | | | 11 | 20.5339 | | |
| Width | 2 | 10.3252 | 9 | 10.2087 | 0.0057 | |
| Depth | 1 | 0.0338 | 8 | 10.1748 | 0.8540 | |
| Rifle | 1 | 3.9872 | 7 | 6.1876 | 0.0458 | |
| Width:Depth | 2 | 1.8577 | 5 | 4.3299 | 0.3950 | |
| Width:Rifle | 2 | 1.9925 | 3 | 2.3374 | 0.3693 | |
| Depth:Rifle | 1 | 0.0406 | 2 | 2.2968 | 0.8404 | |

Advantages of GzLM:

- Not necessary to have a linear relationship, normality or equal variance among groups
- Can model binomial data, good for this data sets because counts vary greatly and are not likely representative of the population (presence/absence more informative)
- Provides probabilities of incidence which can be used to infer distribution

Disadvantages of GzLM

- Does not reflect the actual numbers of fish obtained by sampling
- Cannot easily evaluate model with residual plots, as with GLM diagnostic investigations have to rely on other methods (Venables and Dichmont, 2004).

Appendix C

Age 1 Greenland Cod Counts Suzanne Thompson

Research Question: Does the number of Greenland Cod (*Gadus ogac*) depend on amount of eelgrass at the site of sampling?

Data: Counts of Age 1 Greenland Cod from beach seine tows at several sites conducted at Newman Sound, Terra Nova National Park. Seine tows were taken bi-weekly during from 2006 field season from May till November. Data was collected by the MUN Cod research group, and was retrieved from Bob Gregory at Department of Fisheries and Oceans.

Model 1: General Linear Model

Formal Model: $N = \beta_0 + \beta_E * E + res$

Table C1: ANOVA

| | Sum Sq | Df | F value | Pr(>F) |
|-----------|--------|-----|---------|---------|
| E | 336.2 | 2 | 3.3667 | 0.03724 |
| Residuals | 7189.0 | 144 | | |

Model 2: Generalized Linear Model; error = Negative Binomial

Table C2: LR statistics for Type 1 Analysis

| Source | DF | F-Value | Chi-Square | Pr>Chi-sq |
|--------|----|---------|------------|-----------|
| Е | 2 | 2.75 | 5.50 | 0.0640 |

Table C3: LR statistics for Type 2 Analysis

| Source | DF | F-Value | Chi-Square | Pr>Chi-Sq |
|--------|----|---------|------------|-----------|
| E | 2 | 0.00 | 7.54 | 0.0231 |

Advantages of GzLM:

- Negative Binomial distribution can deal with very over-dispersed abundance data, such is the case here, by adding an extra parameter.
- Shows slightly more sensitive test by giving a lower p-value for Type 2 analysis.

Disadvantages of GzLM:

• Extra parameter (dispersion) makes it difficult to interpret diagnostic tests (not shown) such as normal probability tests, and homogeneity.

Appendix D

Juvenile Fish Counts Suzanne Thompson

Research Question: Does the number of juvenile fish depend on geology configuration of sea bed and time of day?

Data: Counts of juvenile fish (mostly Haddock) from two 5 km transect lines on Western Bank of the Eastern Scotian Shelf. This data was collected by Tow Cam, a video camera which is towed from the back of a ship. Side scan sonar data was also collected of the sea bed configuration, classified in several categories, and matched up with camera data. Data was collected by Department of Fisheries and Oceans, and was retrieved from Bob Gregory.

Model 1: General Linear Model

Formal Model: $N = \beta_0 + \beta_{Geol} * Geol + \beta_{Time} * Time + res$

Table D1: ANOVA

| Source | DF | Type III SS | Mean Square | F Value | Pr>F |
|--------|----|-------------|-------------|---------|---------|
| Geol | 7 | 48.4773 | 6.9253 | 42.72 | <0.0001 |
| Time | 1 | 2.8739 | 2.8739 | 17.73 | <0.0001 |

Model 2: Generalized Linear Model; error = Binomial

Table D2: LR statistics for Type 1 Analysis

| Source | Deviance | DF | Chi-Square | Pr>ChiSq |
|-----------|-----------|----|------------|----------|
| Intercept | 6180.1816 | | | |
| Geol | 5860.7547 | 7 | 319.43 | <0.0001 |
| Time | 5842.8680 | 1 | 17.89 | <0.0001 |

Table D3: LR statistics for Type 2 Analysis

| Source | DF | Chi-Square | Pr>Chi-Square |
|--------|----|------------|---------------|
| Geol | 7 | 308.19 | <0.0001 |
| Time | 1 | 17.89 | <0.0001 |

Advantages of GzLM:

- Because the data has many zeros, analysis can focus less on the abundance, and more on the probability of occurrence.
- Gives a good idea of the magnitude of the differences, because it is based on odds ratios and not on abundance data.

Appendix E

Seed Removal Data Andrew Trant

Research Question: Does seed removal by introduced red fire ant (*Solenopsis invicta* Burren) differ within or between plant species?

Data: These data were collected in South Carolina, 1998. The study area was an abandoned field which had natural populations of the red fire ant. Cages were set up to exclude all known seed predators other than the red fire ant. At each ant mount, six cages were established with seeds placed within cages on platforms. The number of seeds removed was recorded every week for a total of six weeks. The following analysis uses the sums of seed removed over this entire period. These data were published by Seaman and Marino (2003).

Model 1: General Linear Model; log transformed mean counts of seeds removed

Formal Model: $C = \beta_0 + \beta_S *S + \beta_M *M + res$

Table E1: ANOVA

| | Df | Sum Sq | Mean Sq | F-value | Pr(>F) |
|-----------|----|--------|---------|---------|-----------|
| Species | 4 | 32343 | 8086 | 18.111 | 9.273e-11 |
| Mound | 1 | 971 | 971 | 2.1746 | 0.1440 |
| Residuals | 84 | 37501 | 446 | | |

Table E2: General Linear Model

| | Estimate | Std. Error | t-value | Pr(>ltl) |
|-----------------|----------|------------|---------|-----------|
| (Intercept) | 1.182 | 0.336 | 3.522 | 0.000696 |
| SpeciesAmbrosia | 0.649 | 0.372 | 1.746 | 0.084425. |
| SpeciesChenopod | -0.289 | 0.372 | -0.778 | 0.438469 |
| SpeciesPoa | -1.145 | 0.372 | -3.081 | 0.002788 |
| SpeciesSolidago | 0.403 | 0.372 | 1.084 | 0.281360 |
| Mound | -0.007 | 0.022 | -0.318 | 0.751622 |

Model 2: Generalized Linear Model; error = Negative Binomial

Table E3: GzLM summary

| | Estimate | Std. Error | t-value | Pr(>ltl) |
|-----------------|----------|------------|---------|----------|
| (Intercept) | 2.796 | 0.227 | 12.317 | < 2e-16 |
| SpeciesAmbrosia | 1.277 | 0.248 | 5.154 | 1.66e-06 |
| SpeciesChenopod | 0.287 | 0.251 | 1.145 | 0.255 |
| SpeciesPoa | -0.271 | 0.254 | -1.064 | 0.290 |
| SpeciesSolidago | 0.727 | 0.249 | 2.921 | 0.004 |
| Mound | 0.012 | 0.014 | 0.835 | 0.406 |

Table E4: GzLM Chi-square test

| | Df | Deviance | Resid.Df | ResidDev | P(> Chil) |
|---------|----|----------|----------|----------|-----------|
| NULL | | | 89 | 192.926 | |
| Species | 4 | 53.141 | 85 | 139.786 | 1.774e-10 |
| Mound | 1 | 0.763 | 84 | 139.023 | 0.390 |

Advantages of GzLM:

- do not need to transform response variable to fit assumptions of normality
- unlike Poisson models that require and mean to variance ratio of 1, negative binomials models do not
- greater sensitive to p-values compared to similar analysis with general linear model

Disadvantages of GzLM

- not ideal for small sample sizes
- requires estimate of additional parameter (theta)

Appendix F

Oak Recruitment Data Andrew Trant

Research Question: Does the occurrence of fire and other measures of forest structure influence of the distribution of Red Oak (*Quercus rubra*)?

Data: These data were collected November 2007 by Andrew Trant, Ian Morrison and Krista Chin in Kejimkujik National Park, NS as part of a study looking at the role of fire in regeneration and maintenance of Red Oak (*Quercus rubra*) stands. 20 plots were established within the boundary of Kejimkujik National Park in areas that with a known fire history. In each plot, demography and structural data were collected. These data are unpublished.

Model 1: General Linear Model; log transformed mean seedling counts

Formal Model: $C = \beta_0 + \beta_F *F + \beta_{BA} *BA + \beta_{DWD} *DWD + res$

Table F1: ANOVA

| | Df | Sum Sq | Mean Sq | F-value | Pr(>F) |
|------------|----|--------|---------|---------|---------|
| fire | 1 | 4.025 | 4.025 | 1.2219 | 0.2853 |
| basal.area | 1 | 10.512 | 10.512 | 3.1911 | 0.0930. |
| dwdall | 1 | 1.606 | 1.606 | 0.4876 | 0.4950 |
| Residuals | 16 | 52.709 | 3.294 | | |

Model 2: Generalized Linear Model; error = Negative Binomial

Table F2: GzLM summary

| | Estimate | Std. Error | t-value | Pr(>ltl) |
|-------------|----------|------------|---------|----------|
| (Intercept) | -7.287 | 3.140 | -2.320 | 0.034 |

| fire | 3.789 | 1.058 | 3.581 | 0.002 |
|------------|--------|-------|--------|--------|
| basal.area | 0.270 | 0.081 | 3.313 | 0.004 |
| dwdall | -0.037 | 0.018 | -2.085 | 0.053. |

Table F3: GzLM Chi-square test

| | Df | Deviance Resid. | Df | Resid Dev | P(> Chil) |
|------------|----|-----------------|----|-----------|-----------|
| NULL | | | 19 | 216.835 | |
| fire | 1 | 15.823 | 18 | 201.011 | 0.069 |
| basal.area | 1 | 64.206 | 17 | 136.805 | 0.0002476 |
| dwdall | 1 | 13.803 | 16 | 123.002 | 0.089 |

Advantages of GzLM:

- do not need to transform response variable to fit assumptions of normality
- unlike Poisson models that require and mean to variance ratio of 1, negative binomials models do not
- greater sensitive to p-values compared to similar analysis with general linear model

Disadvantages of GzLM

- not ideal for small sample sizes
- requires estimate of additional parameter (theta)

Appendix G

Seedling Herbivory Julia Wheeler

Research question: Does black spruce seedling herbivory differ on different groundcovers and with different levels of slug protection?

Data: Record of herbivory-induced mortality for black spruce seedlings planted on *Cladina, Pleurozium* and simulated bear digs that were either protected from slug herbivory (pesticides) or unprotected. Data are grouped by treatment type (groundcover + level of slug protection) The data collected show the number of herbivory events out of 300 seedlings planted above the treeline of the Mealy Mountains, central Labrador; all data were collected by in the summer of 2007 by Julia Wheeler.

Model 1: General linear model, percent herbivory per plot type

Formal model: $H_{seedling} = \beta_0 + \beta_g * X_g + \beta_s * X_s + residual$

Table G1: ANOVA

 Source
 DF
 Seq SS
 Adj SS
 Adj MS
 F
 P

 ground
 2
 0.050533
 0.050533
 0.025267
 126.33
 0.008

 slug
 1
 0.018150
 0.018150
 0.018150
 90.75
 0.011

 Error
 2
 0.000400
 0.000400
 0.000200
 Total
 5
 0.069083

Model 2: Generalized linear model; error = binomial

Formal model: Odds ($H_{seedling}$) = $e^{(Bref)} e^{(Bg)} e^{(Bs)}$ + error

Table G2: GzLM summary

| Parameter | DF | Estimate | Standard Error | Wald | 95% Confidence Limits | Chi- Square | P >ChiSq |
|---------------|----|----------|-------------------|---------|--------------------------|----------------|----------|
| Intercept | 1 | -0.8350 | 0.2413 | -1.3080 | -0.3621 | 11.97 | 0.0005 |
| Groundcover L | 1 | -0.7580 | 0.3193 | -1.3838 | -0.1323 | 5.64 | 0.0176 |
| Groundcover M | 1 | -1.2063 | 0.3501 | -1.8926 | -0.5200 | 11.87 | 0.0006 |
| Slug O | 1 | 0.6239 | 0.2740 | 0.0868 | 1.1610 | 5.18 | 0.0228 |
| Scale | 0 | 1.0000 | 0.0000 | 1.0000 | 1.0000 | | |

Table G3: GzLM Chi-square test

| Source | Deviance | DF | Square | Pr > ChiSq |
|-------------|----------|----|--------|------------|
| Intercept | 19.5553 | | | |
| Groundcover | 5.4940 | 2 | 14.06 | 0.0009 |
| Slug | 0.2048 | 1 | 5.29 | 0.0215 |

Advantages of GzLM:

- GzLM can consider interactive effects of groundcover and slug controls, while the GLM has insufficient degrees of freedom (analysis performed, but not shown; models above shows outputs when interactive effects are removed from both models)
- Magnitude of the effect can be calculated using odds ratios derived from the GzLM; GLM can only demonstrate a significant difference between herbivory across treatments
- Provides lower p-value for groundcover, the variable of highest interest, indicating a greater ability to detect an effect

Disadvantages of GzLM:

• Less sensitive than ANOVA for detecting effect for the second variable (level of slug control)

Appendix H

Seed predation data Julia Wheeler

Research question: Is conifer seed predation affected by substrate types and presence of slug pesticides?

Data: Record of seed predation for black/white spruce seeds from 93 seed cards distributed across *Cladina, Pleurozium* and simulated bear digs plots that were either treated or untreated with slug pesticide. Data are grouped by treatment type. The cards were collected and seed predation was determined by proportion of seeds removed (out of 10 per card). Seed cards were distributed above the treeline of the Mealy Mountains in central Labrador; all data were collected by in the summer of 2007 by Julia Wheeler.

Model 1: General linear model, percent seed predation per treatment type

Formal model: $H_{seed} = \beta_0 + \beta_{groundcover} * X_{groundcover} + \beta_{slug} * X_{slug} + residual$

Table H1: ANOVA

| Source | DF | Seq SS | Adj SS | Adj MS | F | Р |
|--------|----|----------|----------|----------|-------|-------|
| ground | 2 | 0.041200 | 0.041200 | 0.020600 | 11.04 | 0.083 |
| slug | 1 | 0.004267 | 0.004267 | 0.004267 | 2.29 | 0.270 |
| Error | 2 | 0.003733 | 0.003733 | 0.001867 | | |
| Total | 5 | 0.049200 | | | | |

Model 2: Generalized linear model; error = binomial

Formal model: Odds (H_{seed}) = $e^{(Bref)} e^{(Bgroundcover)} e^{(Bslug)} + error$

Table H2: GzLM summary

| _ | | | Standard | | 6 Confidenc | | |
|-------------|-----|----------|----------|---------|-------------|--------|------------|
| Parameter | DF | Estimate | Error | Lim | nits | Square | Pr > ChiSq |
| Intercept | 1 | 0.5379 | 0.1162 | 0.3102 | 0.7656 | 21.43 | <.0001 |
| Groundcover | L 1 | -0.5437 | 0.1622 | -0.8616 | -0.2257 | 11.23 | 0.0008 |
| Groundcover | M 1 | -0.8252 | 0.1650 | -1.1486 | -0.5017 | 25.00 | <.0001 |
| Slug | 0 1 | 0.1719 | 0.1347 | -0.0921 | 0.4360 | 1.63 | 0.2018 |
| Scale | 0 | 1.0000 | 0.0000 | 1.0000 | 1.0000 | | |

Table H3: GzLM Chi-square test

| | | | Chi- | |
|-------------|----------|----|--------|------------|
| Source | Deviance | DF | Square | Pr > ChiSq |
| Intercept | 31.6419 | | | |
| Groundcover | 3.8410 | 2 | 27.80 | <.0001 |
| Slug | 2.2092 | 1 | 1.63 | 0.2015 |

Advantages of GzLM:

- Again, GzLM can consider interactive effects of groundcover and slug controls, while the GLM has insufficient degrees of freedom
- Detects an effect from the groundcover parameter, where p-value is nonsignificant using ANOVA

Disadvantages of GzLM:

• Residual plots are difficult to interpret for overdispersion

Appendix I

Avian Population Monitoring Paul Regular

Research question: Do murre populations exhibit any long-term trends?

Data: The following long-term count data was collected from census plots of common murres (*Uria aalge*) at Cape St. Mary's, Newfoundland, 1980-2006 (Regular unpublished data) and thick-billed murres (*Uria lomvia*) at Coats Island, Nunavut, 1985-2000²¹⁵. Murres were monitored using similar methods to the Type II approach described by Birkhead and Nettleship²⁰². This technique has been adopted and routinely used to monitor common and thick-billed murre populations throughout the northern hemisphere. For the purposes of statistical analyses of these data, one may construct a GLM (linear regression) with the following formula:

 $C = \beta o + \beta_t * t + Normal \ error$

Where C is the total number of murres attending plots on a particular day of the year when counts were conducted (response variable), and *t* represents the explanatory variable year.

Model 1: General linear model: Murre plot count data

| Summary results | | | | | |
|-----------------|-----|----------------|---------------------|---------|----------|
| | | Estimate | Standard error | t-value | Р |
| (Intercept) | | -4271.8617 | 949.2923 | -4.500 | 9.47e-06 |
| Т | | 2.2902 | 0.4776 | 4.796 | 2.47e-06 |
| ANOVA table | | | | | |
| | df | Sum of Squares | Mean Sum of Squares | F value | Р |
| Т | 1 | 51711 | 51711 | 22.997 | 2.5e-06 |
| Residuals | 326 | 733031 | 2249 | | |

Table I1: ANOVA for Cape St. Mary's plot counts

Table I2: ANOVA for Coats Island plot counts

| Summary results | | | | | |
|-----------------|-----|----------------|---------------------|---------|----------|
| | | Estimate | Standard error | t-value | Р |
| (Intercept) | | -72557.564 | 4398.959 | -16.49 | <2e-16 |
| Т | | 37.229 | 2.207 | 16.87 | <2e-16 |
| ANOVA table | | | | | |
| | df | Sum of Squares | Mean Sum of Squares | F value | Р |
| Т | 1 | 14319562 | 14319562 | 284.56 | <2.2e-16 |
| Residuals | 676 | 34017114 | 50321 | | |

Model 2: Generalized linear model; error = Poisson, negative binomial

Table I3: GzLM summary and Chi-square test for Cape St. Mary's plot counts (error = Poisson)

| Summary results | | | | | |
|------------------|----|------------|----------------|-------------------|-----------|
| | | Estimate | Standard error | z-value | Р |
| (Intercept) | | -1.024e+01 | 1.169e+00 | -8.756 | <2e-16 |
| Т | | 7.987e-03 | 5.883e-04 | 13.577 | <2e-16 |
| Chi-square table | | | | | |
| | df | Deviance | Residual df | Residual deviance | Р |
| NULL | | | 327 | 2921.83 | |
| Т | 1 | 181.65 | 326 | 2740.18 | 2.116e-41 |

| Summary results | | | | | |
|------------------|----|------------|----------------|-------------------|--------|
| | | Estimate | Standard error | z-value | Р |
| (Intercept) | | -3.788e+01 | 4.863e-01 | -77.90 | <2e-16 |
| Т | | 2.272e-02 | 2.439e-04 | 93.14 | <2e-16 |
| Chi-square table | | | | | |
| | df | Deviance | Residual df | Residual deviance | Р |
| NULL | | | 677 | 30659 | |
| Т | 1 | 8720 | 676 | 21939 | 0 |

 Table I4: GzLM summary and Chi-square test Parameter for Coats Island plot counts (error = Poisson)

 Table I5: GzLM summary and Chi-square test for Cape St. Mary's plot counts (error = negative binomial)

| Summary results | | | | | |
|------------------|----|-----------|----------------|-------------------|-----------|
| | | Estimate | Standard error | z-value | Р |
| (Intercept) | | -9.378271 | 3.551322 | -2.641 | 0.00827 |
| Т | | 0.007553 | 0.001787 | 4.228 | 2.36e-05 |
| Chi-square table | | | | | |
| | df | Deviance | Residual df | Residual deviance | Р |
| NULL | | | 327 | 353.04 | |
| Т | 1 | 19.13 | 326 | 333.91 | 1.222e-05 |

Table I6: GzLM summary and Chi-square test for Coats Island plot counts (error = negative binomial)

| Summary results | | | | | |
|------------------|----|------------|----------------|-------------------|-----------|
| | | Estimate | Standard error | z-value | Р |
| (Intercept) | | -36.124955 | 2.885362 | -12.52 | <2e-16 |
| Т | | 0.021838 | 0.001448 | 15.09 | <2e-16 |
| Chi-square table | | | | | |
| | df | Deviance | Residual df | Residual deviance | Р |
| NULL | | | 677 | 918.76 | |
| Т | 1 | 236.03 | 676 | 682.73 | 2.884e-53 |

Table I7: GzLM summary and Chi-square test for Cape St. Mary's plot counts (GAM)

| Summary results | | | | | |
|-----------------|----|-----|--------------------|------------|--------|
| | df | n | Deviance explained | Chi-square | Р |
| s(<i>t</i>) | 3 | 328 | 16.9% | 512 | <2e-16 |

| Summary results | | | | | |
|-----------------|----|-----|--------------------|------------|--------|
| | df | n | Deviance explained | Chi-square | Р |
| s(<i>t</i>) | 3 | 678 | 35.6% | 11091 | <2e-16 |

Table I8: GzLM summary and Chi-square test Parameter for Coats Island plot counts (GAM)

Advantages of GzLM:

- GzLM model does not assume a normal distribution and allows for non-linearity and non-constant variance structures in the data
- P-values generated by GzLM are lower, indicating stronger statistical power

Disadvantages of GzLM

• GzLMs and GAMs include extra parameters to facilitate model fitting; data may be overfitted

Appendix J

Bacterial abundance in ballast water and port water samples Bei Sun

Research Question: If bacterial abundance depends on the two different water masses and pH?

Data: Sample were collected from both ship ballast tanks and receiving ports in three locations: Great Lakes, East Coast and West Coast of Canada. 10 ships from Great Lakes and 30 ships from west coast and east coast of Canada each were sampled from April to November, 2007. Several port samples were also collected in each location through the sampling season. Epi-fluorescent microscope direct count was used to investigate prokaryotical abundance. Physical parameters of those samples were recorded, named temperature, salinity and pH. The preliminary analysis indicated temperature and salinity were correlated, so samples were divided into two water masses by salinity 15%.

Model 1: General Linear Model

Formal model: $A = \beta_0 + \beta_W \bullet W + \beta_{pH} \bullet pH + \beta_W \beta_{pH} pH \bullet W + \varepsilon$

| GLM | Df | Sum Sq | Mean Sq | F value | Pr(>F) |
|--------------|----|------------|------------|---------|-----------|
| Water masses | 1 | 2.3286e+19 | 2.3286e+19 | 19.568 | 3.996e-05 |
| pН | 1 | 1.3086e+17 | 1.3086e+17 | 0.110 | 0.7413 |
| Interaction | 62 | 7.3779e+19 | 1.1900e+18 | | |

Table J1: ANOVA

Model 2: Generalized Linear Model; error = Gamma

| GzLM | Df | Deviance | Resid. Df | Resid. Dev | F | Pr(>F) |
|--------------|----|----------|-----------|------------|---------|-----------|
| NULL | | | 64 | 36.261 | | |
| Water masses | 1 | 13.693 | 63 | 22.568 | 34.0854 | 2.085e-07 |
| pН | 1 | 0.242 | 62 | 22.326 | 0.6012 | 0.4411 |
| Interaction | 0 | 0 | 62 | 22.326 | | |

Table J2: GzLM Chi-square test

Advantages of GzLM:



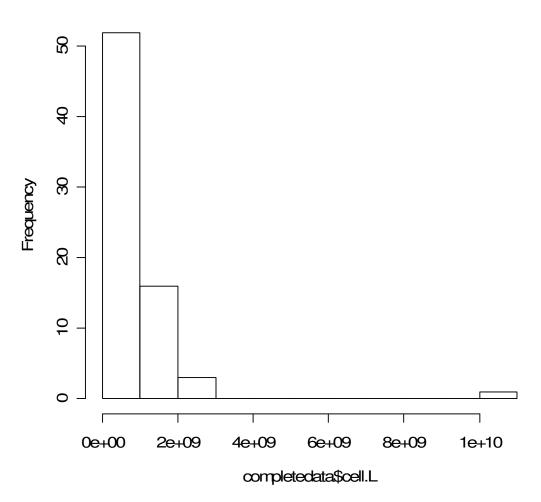


Fig J1: Histogram of bacterial abundance

Advantages of GzLM:

- Gamma can be applied to data with skewed distributions
- Application of the GzLM satisfies the assumption that the fits vs. residuals must be homogenous, and the assumption that the error points must be distributed along a normal line

Disadvantages of GzLM:

• Requires strong knowledge about advanced statistics

Appendix K

Bacterial colonies formed on different quadratic regions of different Petri dishes Bei Sun

Research Question: Is the number of bacterial colonies different on different quadratic regions of a series of Petri dishes?

Data: We counted the number of bacteria colonies in 16 small quadratic regions of the same area in the central part of 5 different Petri dishes.

Model 1: General Linear Model

Formal Model: Two-way ANOVA $C = \beta_0 + \beta_P \bullet P + \beta_r \bullet R + \varepsilon$

Table K1: ANOVA

| GLM | Df | Sum Sq | Mean Sq | F value | Pr(>F) |
|-----------|----|---------|---------|---------|--------|
| Dishes | 1 | 3.025 | 3.025 | 1.3580 | 0.2475 |
| Quadratic | 1 | 3.442 | 3.442 | 1.5454 | 0.2176 |
| Residuals | 77 | 171.520 | 2.228 | | |

Model 2: Generalized Linear Model; error = Poisson

Table K2: GzLM Chi-square test

| GzLM | Df | Deviance | Resid. Df | Resid. Dev | F | Pr(>F) |
|-----------|----|----------|-----------|------------|--------|--------|
| NULL | 79 | 76.092 | | | | |
| Dishes | 1 | 1.205 | 78 | 74.887 | 1.2048 | 0.2724 |
| Quadratic | 1 | 1.371 | 77 | 73.516 | 1.3711 | 0.2416 |

Advantages of GzLM:

• GzLM generates more accurate p-values since it represents the true distribution of the data

Disadvantages of GzLM:

• Requires strong knowledge about advanced statistics

Data set 2 from P395-396 Exercise 8.6

Chapter 8 The Poisson distribution. In Statistics with Applications in Biology and Geology. Preben Blasild, Jorgen Granfeldt (2003) Chapman & hall/CRC