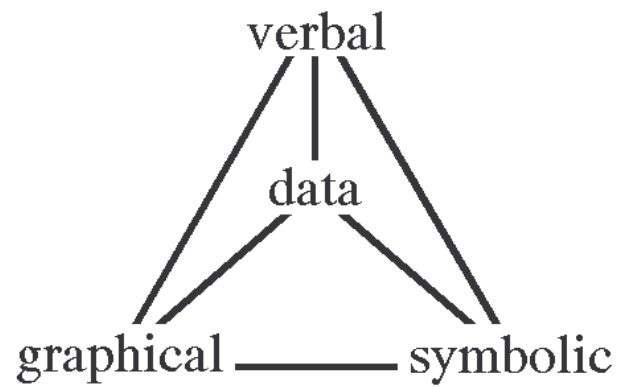


# Handouts in Quantitative Biology

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## Part I Units and Dimensions

**Table 1.** Base and supplementary units in the SI system.

Quantity	Unit	Abbreviation
Length	metre	m
Mass	kilogram	kg
Time	second	s
Thermodynamic temperature	kelvin	K
Amount of substance	mole	mol
Luminous intensity	candela	cd
Electrical current	ampere	A
Planar angle	radian	rad
Solid angle	steradian	sr

**Table 2.** Standard multiples of ratio scale units.

Name	Multiple	Abbreviation	Example
pico	$10^{-12}$	p	pW
nano	$10^{-9}$	n	nW
micro	$10^{-6}$	$\mu$	$\mu$ W
milli	$10^{-3}$	m	mW
centi	$10^{-2}$	c	cW
deci	$10^{-1}$	d	dW
	$10^0$		W
deca	$10^1$	da	daW
hecto	$10^2$	h	hW
kilo	$10^3$	k	kW
mega	$10^6$	M	MW
giga	$10^9$	G	GW

**Table 3.** Units that commonly occur in biology.

Quantity		Unit Name	Unit Symbol	Equivalent Units
Acceleration	angular			$\text{rad}\cdot\text{s}^{-2}$
	linear			$\text{m}\cdot\text{s}^{-2}$
Area		square metre	$\text{m}^2$	
		hectare	ha	$10^4\cdot\text{m}^2$
Concentration				$\text{mol}\cdot\text{m}^{-3}$
Energy (work)		joule	J	$\text{N}\cdot\text{m}$
		kilocalorie	kcal	$4185\cdot\text{J}$
Energy flux				$\text{J}\cdot\text{m}^{-2}\cdot\text{s}^{-1}$
Force		newton	N	$\text{kg}\cdot\text{m}\cdot\text{s}^{-2}$
Frequency		hertz	Hz	$\text{s}^{-1}$
Light	Luminance			$\text{cd}\cdot\text{m}^{-2}$
	Luminous flux	lumen	lm	$\text{cd}\cdot\text{sr}$
	Illuminance	lux	lx	$\text{lm}\cdot\text{m}^{-2}$
		footcandle	fc	$10.764\cdot\text{lx}$
	Photon flux	einstein	E	1·mole
Mass density				$\text{kg}\cdot\text{m}^{-3}$
Mass flow				$\text{kg}\cdot\text{s}^{-1}$
Mass flux				$\text{kg}\cdot\text{m}^{-2}\cdot\text{s}^{-1}$
Power		watt	W	$\text{J}\cdot\text{s}^{-1}$
Pressure (stress)		pascal	Pa	$\text{N}\cdot\text{m}^{-2}$
Surface tension				$\text{N}\cdot\text{m}^{-1}$
Velocity	angular			$\text{rad}\cdot\text{s}^{-1}$
	linear			$\text{m}\cdot\text{s}^{-1}$
Viscosity	dynamic			$\text{Pa}\cdot\text{s}$
	kinematic			$\text{m}^2\cdot\text{s}^{-1}$
Volume		cubic metre	$\text{m}^3$	
		litre	l	$10^{-3}\text{m}^3$
Volume flow rate				$\text{m}^3\cdot\text{s}^{-1}$
Wavelength				m
Wavenumber				$\text{m}^{-1}$

**Table 4.** Rules for working with dimensions.

From D.S. Riggs (1963) *The Mathematical Approach to Physiological Problems*. MIT Press.

---

1. All terms in equation must have the same dimensions.  
Terms separated by + - or = .
  2. Multiplication and division must be consistent with rule 1.
  3. Dimensions are independent of magnitude.  
 $dx/dt$  is the ratio of infinitesimals,  
but still has dimensions of  $x/t = \text{Length/Time}$ .
  4. Pure numbers (e,  $\pi$ ) have no dimensions.  
Exponents and percentages have no dimensions.
  5. Multiplication by a dimensionless number does not change dimensions.
- 

Working with Dimensions--Examples.

1. According to Holligan et al 1984 (*Marine Ecology Progress Series* 17:201) the vertical flux of nutrients through the ocean's thermocline is:

$$F_N = K_v \Delta N / \Delta Z$$

where  $F_N$  is the vertical flux of nutrients (milligram-atoms  $\text{m}^{-2} \text{s}^{-1}$ )

$K_v$  is the vertical eddy diffusivity ( $10^{-4} \text{m}^2 \text{s}^{-1}$ )

$\Delta N$  is the nitrate difference across the thermocline (mg-atoms)

$\Delta Z$  is the thickness of the thermocline (metres)

Write out dimensions beneath each symbol in the equation.

Is this equation dimensionally homogeneous? \_\_\_\_\_

Work out the dimensions of  $\Delta N$  required to make the equation homogeneous \_\_\_\_\_

Work out the units of  $\Delta N$  required to make the equation homogeneous \_\_\_\_\_

$M = \text{Mass}$                        $M L^{-1} = \text{mass gradient}$

$M L^{-2} = \text{mass density}$     $M L^{-3} = \text{mass concentration}$

Based on this,  $\Delta N$  must be the difference in nitrate \_\_\_\_\_ across the thermocline.

More Examples with Units and Dimensions (continued)

2. A series of experimental measurements by Holligan *et al* suggest that the vertical flux of nutrients through the thermocline follows an exponential relation:

$$F_N = \alpha (K_V \Delta N / \Delta Z)^{3/4}$$

What units does  $\alpha$  have? \_\_\_\_\_

What dimensions does  $\alpha$  have? \_\_\_\_\_

3. Another series of experiments by Holligan *et al* suggest that nutrient flux depends upon the temperature gradient across the thermocline.

$$F_N = \beta (\Delta T / \Delta Z)^{-1/3}$$

$$\Delta T / \Delta Z = ^\circ\text{C}/\text{metre}$$

What units does  $\beta$  have? \_\_\_\_\_

What dimensions does  $\beta$  have? \_\_\_\_\_

Elementary statistics courses for biologists tend to lead to the use of a stereotyped set of tests:

- 1 without critical attention to the underlying model involved;
- 2 without due regard to the precise distribution of sampling errors;
- 3 with little concern for the scale of measurement;
- 4 careless of dimensional homogeneity;
- 5 without considering the ideal transformation;
- 6 without any attempt at model simplification;
- 7 with too much emphasis on hypothesis testing and too little emphasis on parameter estimation.

M.J. Crawley. 1993. *GLIM for Ecologists*. (London, Blackwell)

## Euclidean and Fractal Dimensions in Biology -- References

- Gunther, B. 1975. Dimensional analysis and the theory of biological similarity. *Physiological Reviews* 55: 659-698.
- Hastings, H. M. and G. Sugihara. 1993. *Fractals: a User's Guide for the Natural Sciences*. Cambridge University Press.
- Mandelbrot, B.B. 1977. *Fractals: Form, Chance, and Dimension*. San Francisco: Freeman.
- Pennycuik, C.J. *Newton Rules Biology: A Physical Approach to Biological Problems*. Oxford University Press.
- Platt, T.R. and W. Silvert. 1981. Ecology, physiology, allometry, and dimensionality. *Journal of Theoretical Biology* 93: 855-860.
- Schneider, D.C. 1994. *Quantitative Ecology: Spatial and Temporal Scaling*. San Diego: Academic Press.
- Stahl, W.R. 1961, 1962. Dimensional analysis in mathematical biology. *Bulletin of Mathematical Biophysics* 23: 355-376, 24: 81-108.
- Sugihara, G., B. Grenfell, and R.M. May. 1990. Applications of fractals in ecology. *Trends in Resereach in Ecology and Evolution*. 5: 79-87.
- <short, highly readable account, including how to estimate  $km^d$ >
- West, B.J. and A.L. Goldberger. 1987. Physiology in fractal dimensions. *American Scientist* 75: 351-365.

## Part II. The General Linear Model.

### Notation for Frequency Distributions and Probability Functions.

There is no standard notation for frequency distributions and probability functions: the notation will vary from text to text. Here are some notational conventions that tend to be widely used. Equivalent notation is also shown.

An empirical distribution constructed from a sample of size  $n$  can be expressed in any of four different ways:

$F(Q = k)$	histogram of values	frequencies
$F(Q = k)/n$	histogram of proportions	relative frequencies
$F(Q \leq k)$	histogram of cumulative values	cumulative frequencies
$F(Q \leq k)/n$	histogram of proportions	cumulative relative frequencies

Theoretical distributions can be either discrete (binomial, Poisson) or continuous (normal, chisquare,  $F$ ,  $t$ ). These are functional expressions. The probability density function pdf is a function for the probability, or relative frequency. The cumulative density function cdf is for the cumulative probability, or cumulative frequency. These function can thus be considered models for the frequency distribution obtained from data.

	Observed $n = \text{sample}$	Expected $N = \text{population}$	$k$ is discrete $x$ is continuous	$Q$ is measured $X$ is continuous
Frequency	$F(Q = k)$ $n \cdot \Pr(Q \leq k)$ $n \cdot \Pr(X \leq x)$ $N \cdot \Pr(Q \leq k)$ $N \cdot \Pr(X \leq x)$	Frequency of $Q$ in the sample of size $n$ Expected frequency that $Q$ in sample, limited to $k$ values Expected frequency $X$ in sample, $X$ continuous Expected frequency that $Q$ in population, $k$ values only Expected frequency $X$ in population, $X$ continuous		(the histogram)
Relative Frequency	$F(Q = k)/n$ $\Pr(Q = k)$ $\Pr(X = x)$	Proportion of $Q$ in the sample of size $n$ Probability that $Q = k$ Probability that $X = x$		probability mass function, pmf probability density function, pdf
Cumulative Frequency	$F(Q \leq k)$ $n \cdot \Pr(Q \leq k)$ $n \cdot \Pr(X \leq x)$ $N \cdot \Pr(Q \leq k)$ $N \cdot \Pr(X \leq x)$	Cumulative frequency of $Q$ Expected frequency that $Q \leq k$ in sample, limited to $k$ values Expected frequency $X \leq x$ in sample, $X$ continuous Expected frequency that $Q \leq k$ in population, $k$ values only Expected frequency $X \leq x$ in population, $X$ continuous		
Cum. Relative Frequency	$F(Q \leq k)/n$ $\Pr(Q \leq k)$ $\Pr(X \leq x)$	Proportion of $Q \leq k$ in the sample of size $n$ Probability that $Q \leq k$ Probability that $X \leq x$		cumulative mass function, cmf cumulative density function, cdf



Notation for Frequency Distributions and Probability Functions.

Equivalent notation	$\Pr(Q = k)$	$f(x)$	pmf	$P(Q = k)$	for discrete variables
	$\Pr(X = x)$	$f(x)$	pdf	$P(X = x)$	for continuous
	$\Pr(Q \leq k)$	$F(x)$	cmf	$P(Q \leq k)$	for discrete variables
	$\Pr(X \leq x)$	$F(x)$	cdf	$P(X \leq x)$	for continuous

**Table 5.** Key for choosing the frequency distribution of a statistic.

Statistic is the population mean

If data are normal or cluster around a central value

If sample is large ( $n > 30$ ) . . . . . Normal distribution

If sample is small ( $n < 30$ ) . . . . . t distribution

If data are Poisson . . . . . Poisson distribution

If data are Binomial . . . . . Binomial distribution

If data do not cluster around central value, examine residuals (deviations from the mean)

If residuals are normal or cluster around a central value

If sample is large ( $n > 30$ ) . . . . . Normal distribution

If sample is small ( $n < 30$ ) . . . . . t distribution

If residuals are not normal . . . . . Empirical (bootstrap)

Statistic is the population variance

If data are normal or cluster around a central value . . . . . Chi-square

If data do not cluster around central value

If sample is large ( $n > 30$ ) . . . . . Chi-square

If sample is small ( $n < 30$ ) . . . . . Empirical (bootstrap)

Statistic is the ratio of two variances (ANOVA tables)

If data are normal or cluster around a central value . . . . . F-distribution

If data do not cluster around a central value, calculate residuals

If residuals are normal or cluster around a central value . . . . . F-distribution

If residuals do not cluster around central values

If sample is large ( $n > 30$ ) . . . . . F-distribution

If sample is small ( $n < 30$ ) . . . . . Empirical

Statistic is none of the above

Search statistical literature for appropriate distribution  
or confer with statistician

If not in literature or cannot be found . . . . . Empirical

Empirical distributions are generated by taking all permutations, by sampling permutations, or by subsampling (bootstrap methods).

**Table 6.** Generic recipe for calculating a confidence limit.

1. State population; state the statistic of interest.
2. Calculate an estimate of the statistic from data
3. Determine the distribution of the estimate.
4. State tolerance for Type I error.
5. Write a probability statement about the estimate or statistic.
6. Plug values into the statement to obtain confidence limits.
7. Make a statement about the probability that the line  
(or limits) include the true value.  
This statement is not about the statistic or estimate.

Strangely, the motto chosen by the founders of the Statistical Society in 1834 was *Aliis exterendum*, which means "Let others thrash it out." William Cochran confessed that "it is a little embarrassing that statisticians started out by proclaiming what they will not do."

E. A. Gehan and N. A. Lemak. 1995. *Statistics in Medical Research: Developments in Clinical Trials* (Plenum Press).

Fisher's famous paper of 1922, which quantified information almost half a century ago, may be taken as the fountainhead from which developed a flow of statistical papers, soon to become a flood. This flood, as most floods, contains flotsam much of which, unfortunately, has come to rest in many text books. Everyone will have his own pet assortment of flotsam; mine include most of the theory of significance testing, including multiple comparison tests, and non parametric statistics.

John Nelder, Rothamsted Experimental Station. (Fisher's successor as Director of the Statistics Department, and pioneer of generalised linear models). From: *Mathematical Models in Ecology*, British Ecological Society Symposium 1971.

**Table 7.** Generic recipe for decision making with statistics.

- 
- |   |          |
|---|----------|
| 1. State population, conditions for taking sample.  |          |
| 2. State the model or measure of pattern  | ST       |
| 3. State Null Hypothesis about the population   | $H_o$    |
| 4. State Alternative Hypothesis   | $H_a$    |
| 5. State criterion (tolerance) for Type I error   | $\alpha$ |
| 6. State frequency distribution that gives probability of outcomes when the Null Hypothesis is true. Choices are: |          |
| Permutations, i.e. distribution of all possible outcomes when $H_o$ is true;                                      |          |
| Empirical distribution obtained by random sampling of all possible outcomes when $H_o$ is true;                   |          |
| Cumulative distribution function (cdf) that applies when $H_o$ is true;   |          |
| State assumptions when using a cdf such as normal, F, t, or chisquare.  |          |
| 7. Calculate the statistic. This is the observed outcome.   |          |
| 8. Calculate the p-value for the observed outcome relative to distribution of outcomes when $H_o$ is true.        |          |
| 9. If p less than $\alpha$ then reject $H_o$ and accept $H_a$   |          |
| If p greater than $\alpha$ then accept $H_o$ .  |          |
| 10. Report statistic, p-value, sample size.   |          |
| Declare decision.   |          |
- 

Equivalent method (less informative) based on just a statistical table, no computer

8. Calculate outcome corresponding to  $\alpha$
9. If observed outcome  $>$  outcome @  $\alpha$  then reject  $H_o$ , accept  $H_a$ .  
If observed outcome  $\leq$  outcome @  $\alpha$  then accept  $H_o$ .
10. Report statistic, p-value, and sample size. Declare decision.

This latter method is less informative, because the observed p-value does not get reported. This method was made necessary by the cumbersome tables for frequency distribution. With modern computers it is possible to calculate an exact p-value for any statistic. The method of reporting an exact p-value is preferred to the method based on tables.

---

**Table 8** Generic Recipe for data analysis with the General Linear Model.

This is a modification of the Generic Recipe for Hypothesis testing.

The equation links one or more response variables to one or more explanatory variables, via parameters (means and slopes).

For reports, use the methods section to:

state that the residuals were examined for normality, homogeneity, and independence;

**Table 9.** Commonly used tests, based on the General Linear Model.

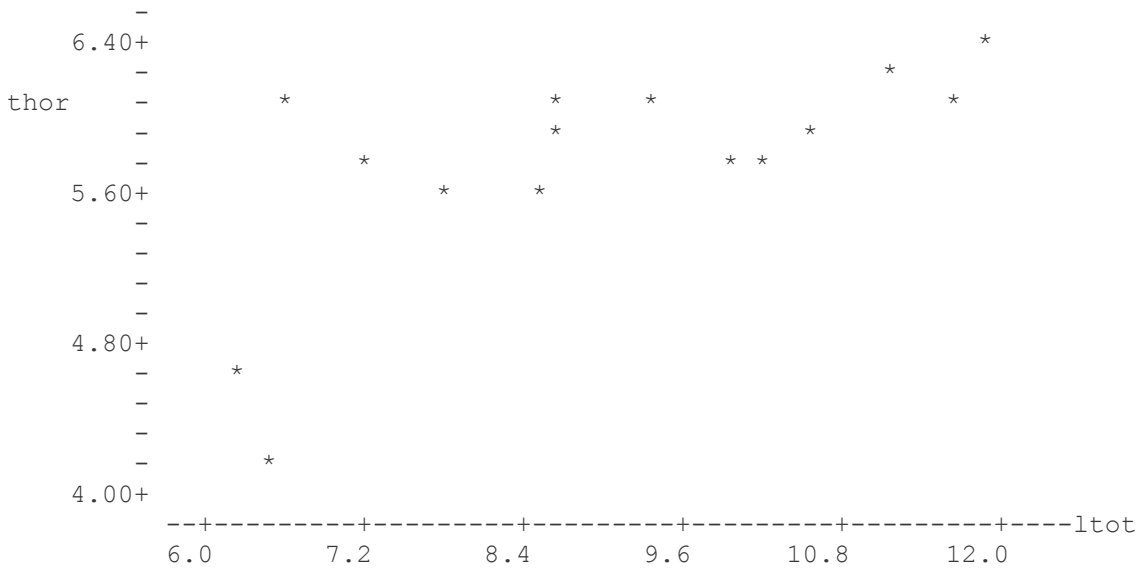
Analysis	Response Variable	Explanatory Variable	Interaction?	Comments
t-test	1 ratio	1 nominal	Absent	compares two means
1-way ANOVA	1 ratio	1 nominal	Absent	compares 3 or more means in 1 category
2-way ANOVA	1 ratio	2 nominal	Present	tests for interactive effects compares means in 2 categories, if no interaction
Paired Comparison	1 ratio	2 nominal	Assumed Absent	compares 2 means in 1 category, controlled for 2nd category (blocks or units)
Randomized Blocks	1 ratio	2 nominal	Assumed Absent	compares 3 or more means in 1 category, controlled for 2nd category (blocks or sampling units)
Hierarchical ANOVA	1 ratio	$\geq 2$ nominal	Absent	nested comparisons of means
ANCOVA	1 ratio	$\geq 1$ ratio $\geq 1$ nominal	Present Absent	compares two or more slopes compares means, controlled for slopes
Regression	1 ratio	1 ratio	Absent	tests linear relation of response to explanatory
Multiple Regression	1 ratio	$\geq$ ratio	Assumed Absent	tests linear relation to 2 explanatory variables relation expressed as a plane

## Correlation (srbx15\_7.out)

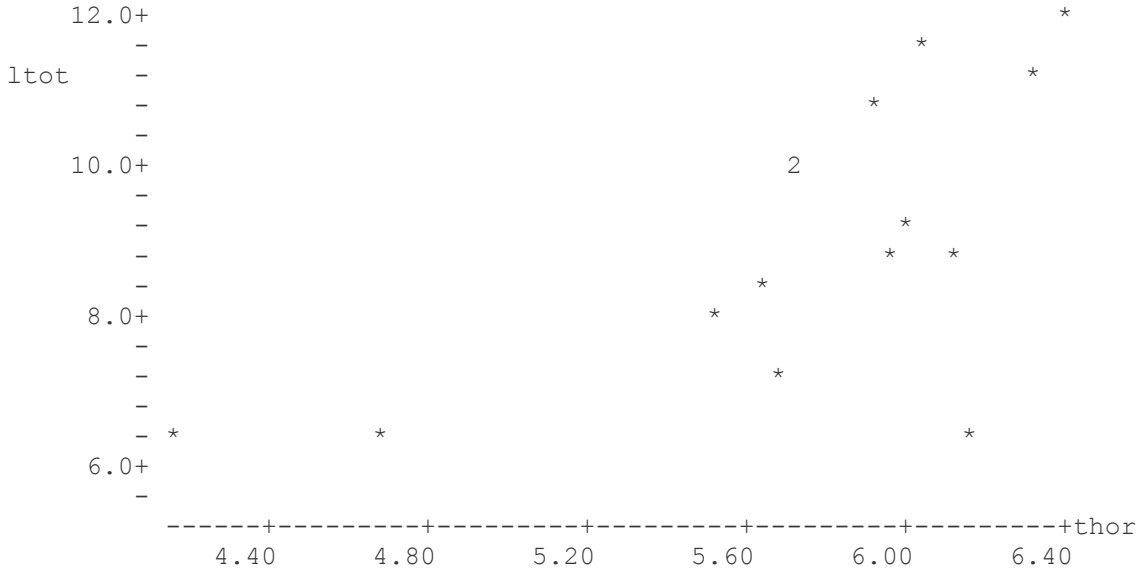
Thorax length data from Box 15.7 in Sokal and Rohlf (1995), p 594.

```
MTB > read 'a:srbx15_7.dat' c1 c2;
SUBC> nobs = 15.
      15 ROWS READ
```

```
MTB > name c1 'ltot' c2 'thor'
MTB > plot c2 c1
```



```
MTB > plot c1 c2
```



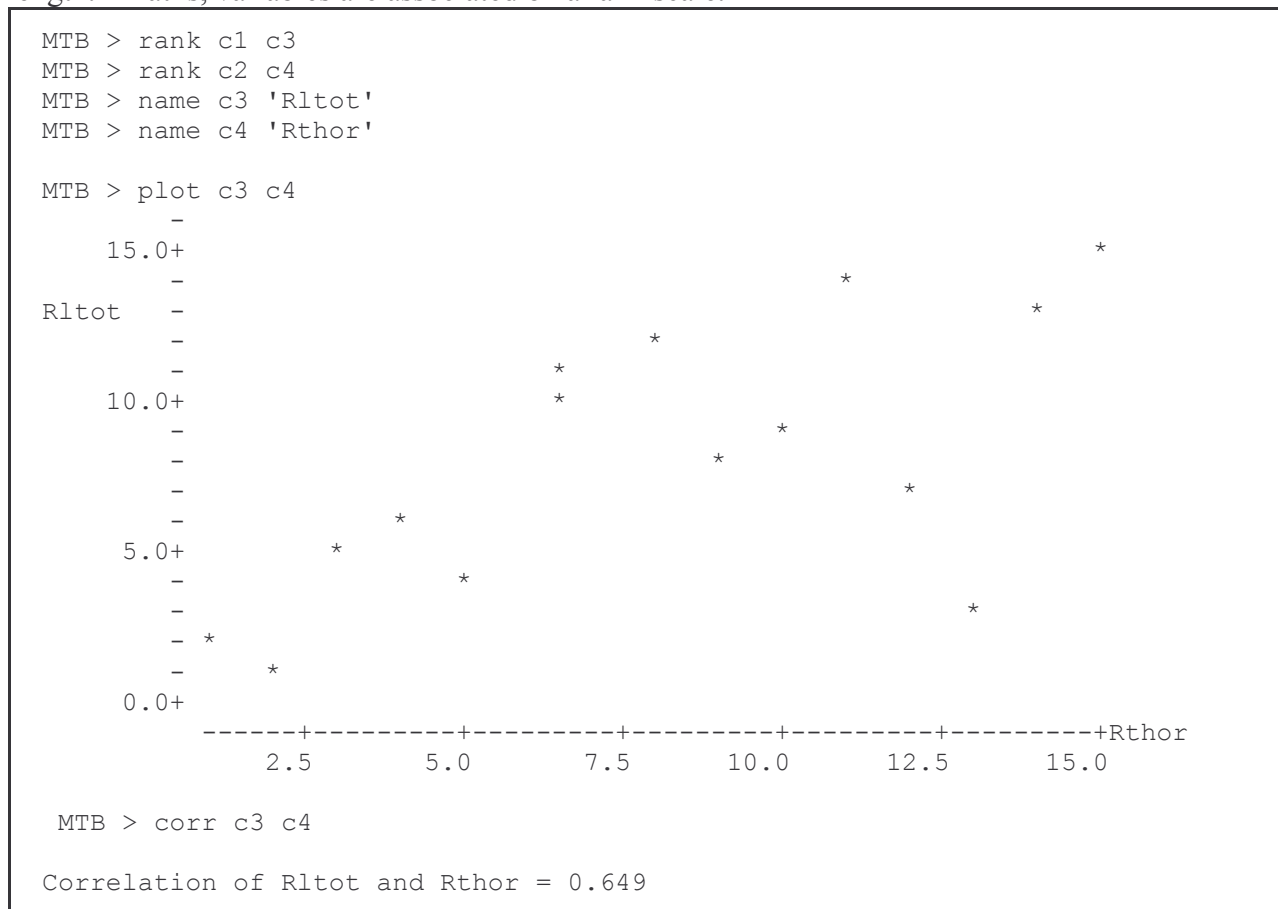
Total length of 15 aphid stem mothers and the mean thorax length of their parthenogenetic offspring.

Judging from these graphs, a linear model of association did not look acceptable. The following models were then investigated by transforming one or both variables, plotting, and examining the plot to see if it was linear (no bowls or arches).

ltot	log(lthor)
log(lot)	lthor
log(ltot)	log(lthor)
ltot	1/lthor
ltot	lthor <sup>3</sup>

The last two were a slight improvement over the first three, but none of the plots could be viewed as linear.

Next, try a model based on monotonic relation: thorax length increases monotonically with total length. That is, variables are associated on a rank scale.



This is called the Spearman Rank correlation coefficient. It is a measure of monotonic relation. It measures the linear relation between the **ranks** of the variables.



How does this measure of monotonic association compare with a measure of linear association?

```
MTB > corr c1 c2 m1
Correlation of ltot and lthor = 0.650
```

This is the Pearson correlation, a measure of the linear association between the variables. In this example, the measure of linear association turns out to be the same as the measure of monotonic association.

So far 6 different models have been tried, none could be considered acceptable, based on lack of bowls or arches in the residuals (deviations from line), as judged by eye. Perhaps the problem is that the data are heterogeneous. There appears to be a positive relation, but some of the data points do not conform to this relation. In particular, it seems that any thorax length is possible at low total lengths ( $ltot < 7$  micrometer units). Let's assume that something different is happening at low total lengths, and just examine the relation between variables when  $ltot > 7$  micrometer units.

```
MTB > let c1(5) = 0/0
MTB > let c1(5) = 0/0
J
*** VALUES OUT OF BOUNDS DURING OPERATION AT J

MTB > let c1(8) = 0/0
MTB > let c1(9) = 0/0
MTB > plot c1 c2
ltot
-
-
-
11.2+
-
-
-
9.6+
-
-
-
8.0+
-
-
-
-----+-----+lthor
5.60 5.76 5.92 6.08 6.24 6.40

N* = 3
```

This looks acceptably linear.

Now compute Pearson correlation, placing the coefficient into k1 for later use.

```
MTB > corr c1 c2 m1

Correlation of ltot and lthor = 0.664

MTB > copy m1 c3 c4
MTB > let k1 = c3(2)
MTB > print k1
K1          0.663741
```

Next compute t-statistic, with  $H_0$  that the true correlation is zero.

```
MTB > let k2 = k1*sqrt((12-2)/(1-k1**2))
MTB > print k2
K2          2.80620
```

Compute p-value from cumulative distribution function, for t distribution.

```
MTB > cdf k2;
SUBC> t 10.
      2.8062      0.9907
MTB > let k3 = (1-.9907)*2
MTB > print k3
K3          0.0186000
```

Note multiplication by 2, the cumulative distribution function yields proportion of outcomes smaller than  $t = 2.8062$ , which comes to 99.07% of the outcomes.

The right tail is thus approximately  $1 - 0.9907 = 0.0093$  and both tails together comes to approximately 1.8% ( $p = 0.0186$  exactly).

Summary.

For non-linear (monotonic) model, use ranks. Compute rank correlation.

For linear model (relation described by straight line) use Pearson correlation.

## Multivariate Analysis -- References

Cooley, W. W. and P. R. Lohnes (1971). *Multivariate Data Analysis*. Wiley & Sons, New York.

Gittens, R. Canonical Analysis. *Biomathematics* **12**. Springer-Verlag, Berlin.

Ludwig, J. A. and J. F. Reynolds (1988). *Statistical Ecology*. Wiley & Sons, New York.

Kim, J. and C. W. Mueller (1978). *Introduction to Factor Analysis. What it is and How to do it*. Sage Publications, London.

Morrison, D. F. (1976). *Multivariate Statistical Methods*. McGraw-Hill, New York.

Pielou, E. C. (1984). *The Interpretation of Ecological Data*. Wiley & Sons, New York.

Seal, H. L. (1964). *Multivariate Statistical Analysis for Biologists*. Methuen, London.

Van de Geer, J. P. (1971). *Introduction to Multivariate Analysis for the Social Sciences*. W. H. Freeman, San Francisco.

Most statistical packages (such as SAS, BMDP, SYSTAT, SPSS) include references.

There are aspects of statistics other than its being intellectually difficult that are barriers to learning. For one thing, statistics does not benefit from a glamorous image that motivates students to persist through tedious and frustrating lessons....there are no TV dramas with a good-looking statistician playing the lead, and few mother's chests swell with pride as they introduce their son or daughter as "the statistician."

C.T. Le and J.R. Boen. 1995. *Health and Numbers: Basic Statistical Methods*. Wiley.

## Autocorrelated Data -- References

Box, G. E. P. and G. H. Jenkins (1976). *Time Series Analysis: Forecasting and Control*. Holden-Day, San Francisco.

<the basic text in time series analysis>

Cressie, N. A. C. (1991). *Statistics for Spatial Data*. John Wiley, New York

<extensive treatment of topic, fairly mathematical>

Diggle, P. J. (1983). *Statistical Analysis of Spatial Point Patterns*. Academic Press, London.

<somewhat mathematical, emphasizes use of randomization tests>

Griffith, D. A. (1987). *Spatial Autocorrelation*. Resource Publications in Geography, American Society of Geographers.

<accessible treatment with examples>

Platt, T. and K. L. Denman (1975). Spectral analysis in ecology. *Annual Review of Ecology and Systematics* 6: 189-210.

<reviews one technique: analysis in the frequency domain>

Ripley, B. D. (1981). *Spatial Statistics*. Academic Press, London.

<comprehensive coverage of topics, fairly mathematical>

Upton, G. J. and B. Fingleton (1985). *Spatial Data Analysis by Example*. Vol. I. Point Pattern and Quantitative Data. John Wiley & Sons, Chichester.

<highly accessible because of examples; short on conceptual linkages>

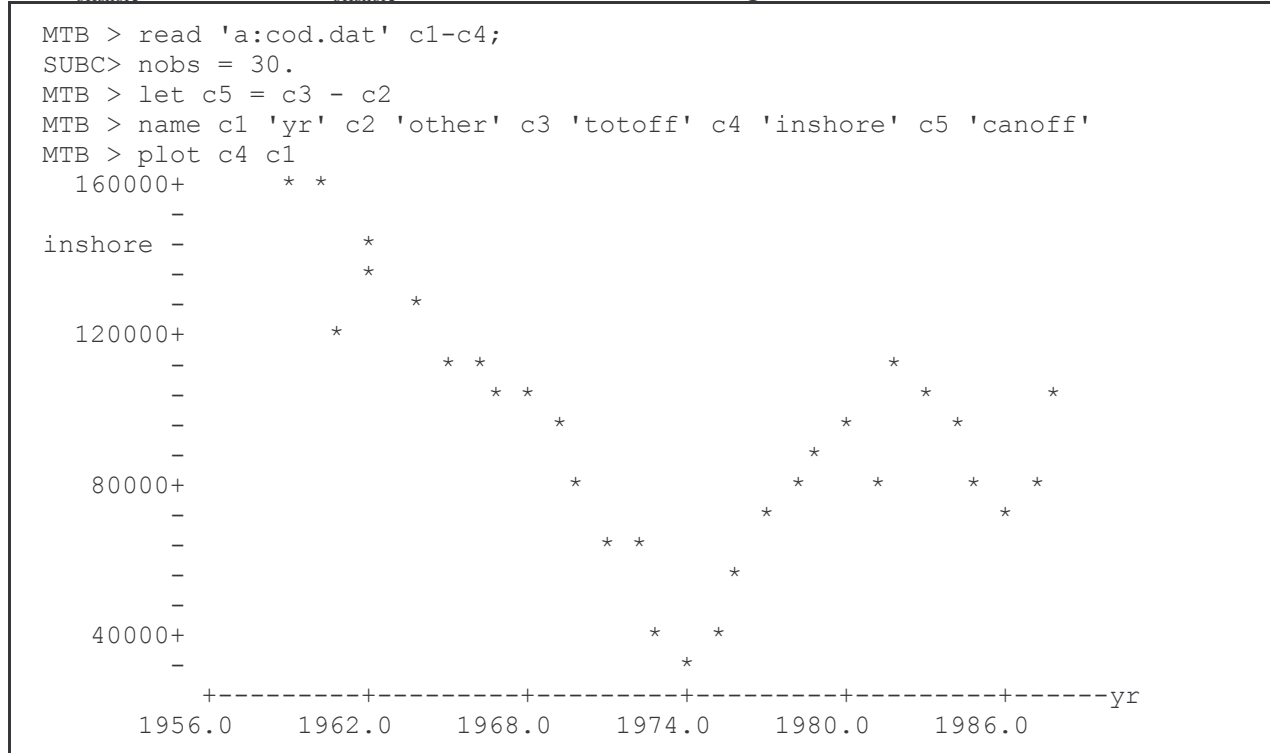
Most statistical packages (such as SAS, BMDP, SYSTAT, SPSS) include references.

## GLM: Autocorrelated Data (codacf.out)

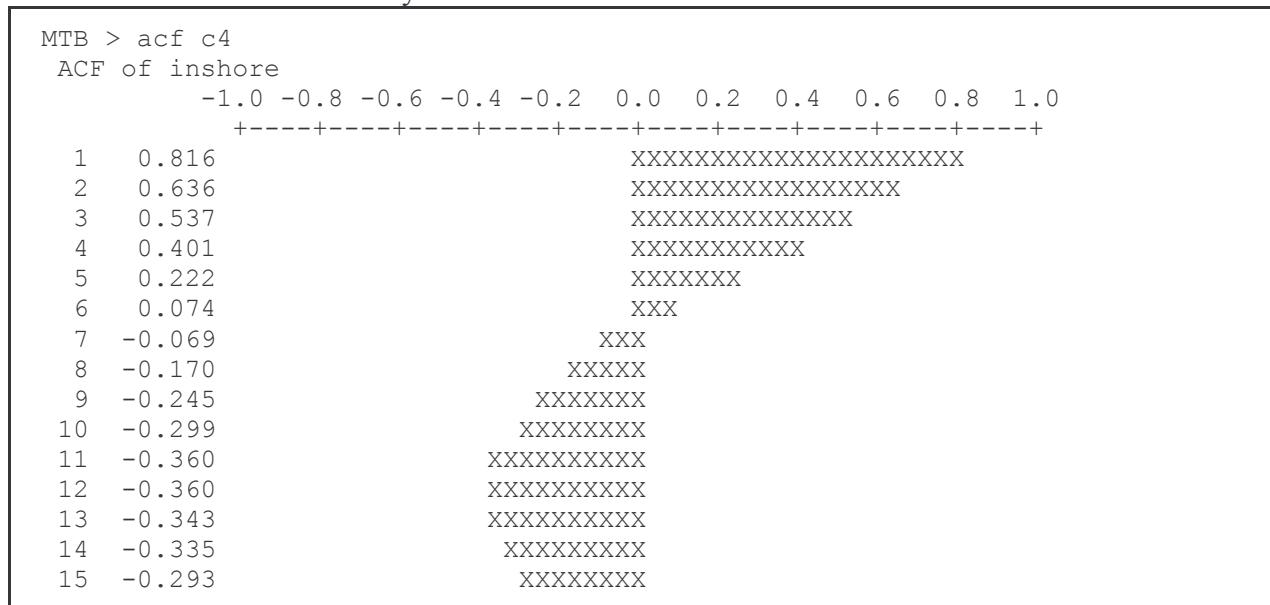
Cod (*Gadus morhua*) catch data.

Catches from the northwest Atlantic, NAFO division 2J3KL are divided into Canadian offshore, other offshore, and inshore.

$\text{Total}_{\text{offshore}} = \text{Other} + \text{Can}_{\text{offshore}}$ . Catches in tonnes =  $10^3$  kg.



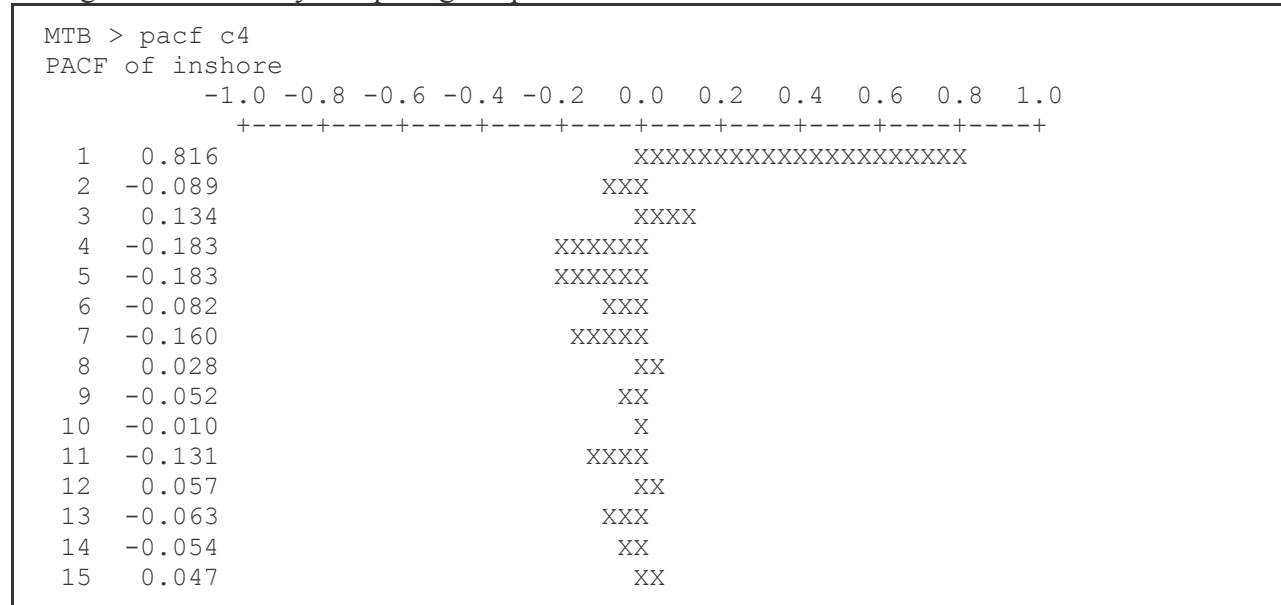
Are the inshore catches serially correlated?



Yes. Inshore catches are strongly correlated.  $r = +0.816$  at lag of 1 year. This means that if catches are high in one year, they will be high the year before or the year after. Catches negatively correlated at lag of 11 years ( $r = -0.36$ ).

What is best model to describe the relation? The two choices are moving average and autoregressive. Moving average means that catch in any one year depends on combined effects of several previous years. Autoregressive means that catch in any one year is related primarily to effects during a fixed time previously.

The shape of the autocorrelation function suggests that this catch is best described as moving average. Check this by computing the partial autocorrelation with PACF command



The shape of the partial autocorrelation function also indicates that catch is related to several prior years (moving average) rather than to year at fixed time in past.

#### Conclusions:

Inshore catches strongly autocorrelated.

A moving average model is best guess for a statistical model.

Next Analysis: Can inshore catches be predicted from offshore catches?

```
MTB > regress c4 1 c5;
SUBC> residuals c8.
```

The regression equation is  
 $\text{inshore} = 95000 - 0.028 \text{ canoff}$

Predictor	Coef	Stdev	t-ratio	p
Constant	95000	7851	12.10	0.000
canoff	-0.0285	0.1338	-0.21	0.833

s = 32914      R-sq = 0.2%      R-sq(adj) = 0.0%

Analysis of Variance

SOURCE	DF	SS	MS	F	p
Regression	1	49014084	49014084	0.05	0.833
Error	28	30333534208	1083340544		
Total	29	30382548992			

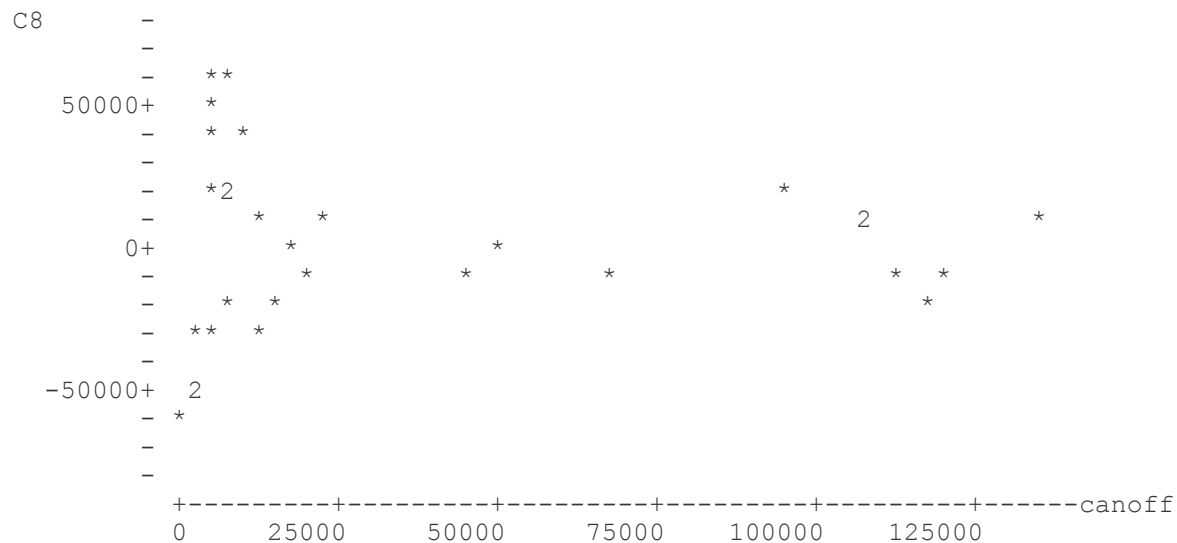
Unusual Observations

Obs.	canoff	inshore	Fit	Stdev.Fit	Residual	St.Resid
1	4515	159492	94871	7477	64621	2.02R

R denotes an obs. with a large st. resid.

Is this model acceptable? Check assumption A, linear relation.

```
MTB > plot c8 c5
```



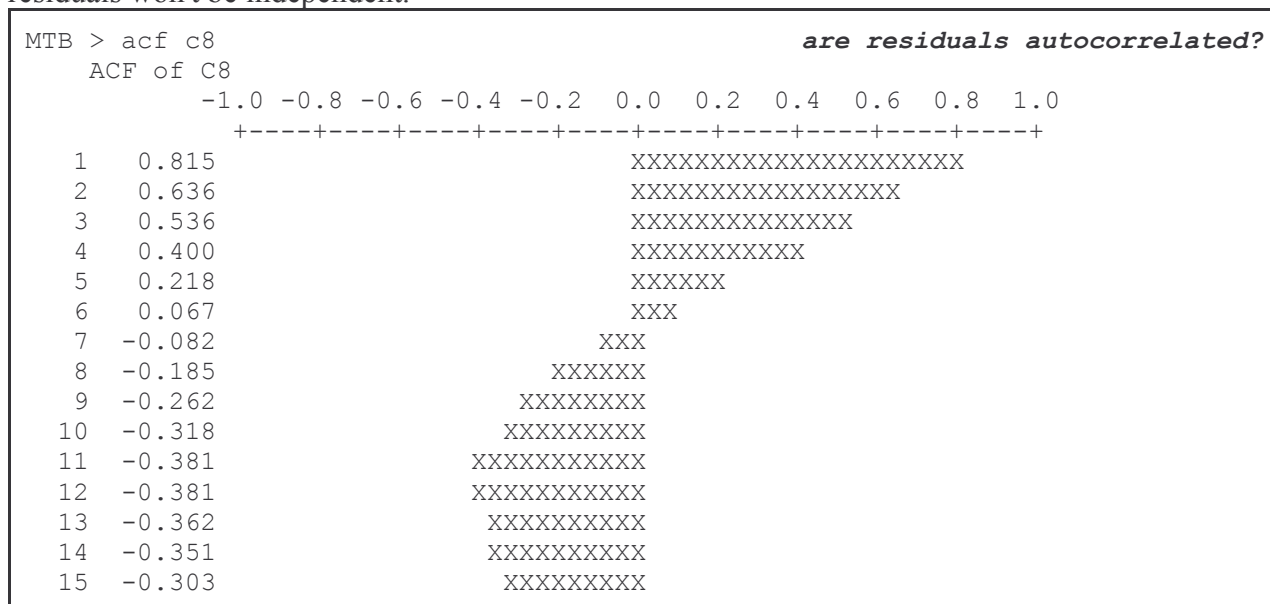
No bowls or arches, so linear model acceptable.

Next, investigate the assumptions concerning errors.

B1  $\text{sum}(\text{errors}) = 0$  ? Yes, because least squares used in regression.

B2 errors independent ?

The catches are strongly autocorrelated, so residuals are also likely to be autocorrelated. If the residuals are autocorrelated, then p-values based on this model will be in error because the residuals won't be independent.



The residuals are not independent. p-value cannot be trusted.

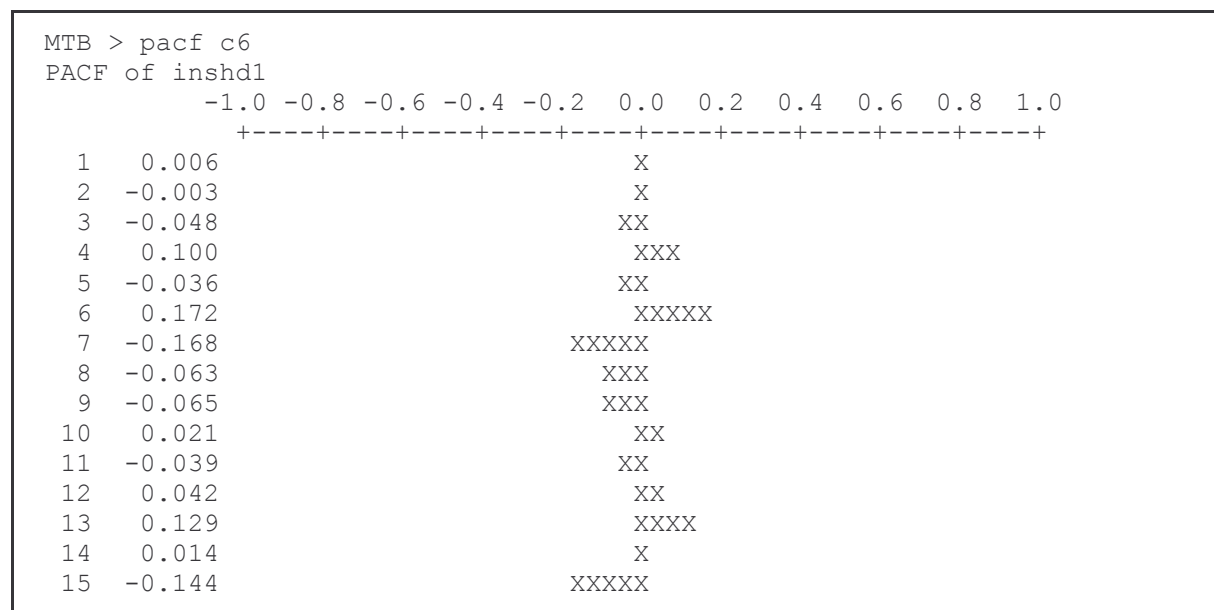
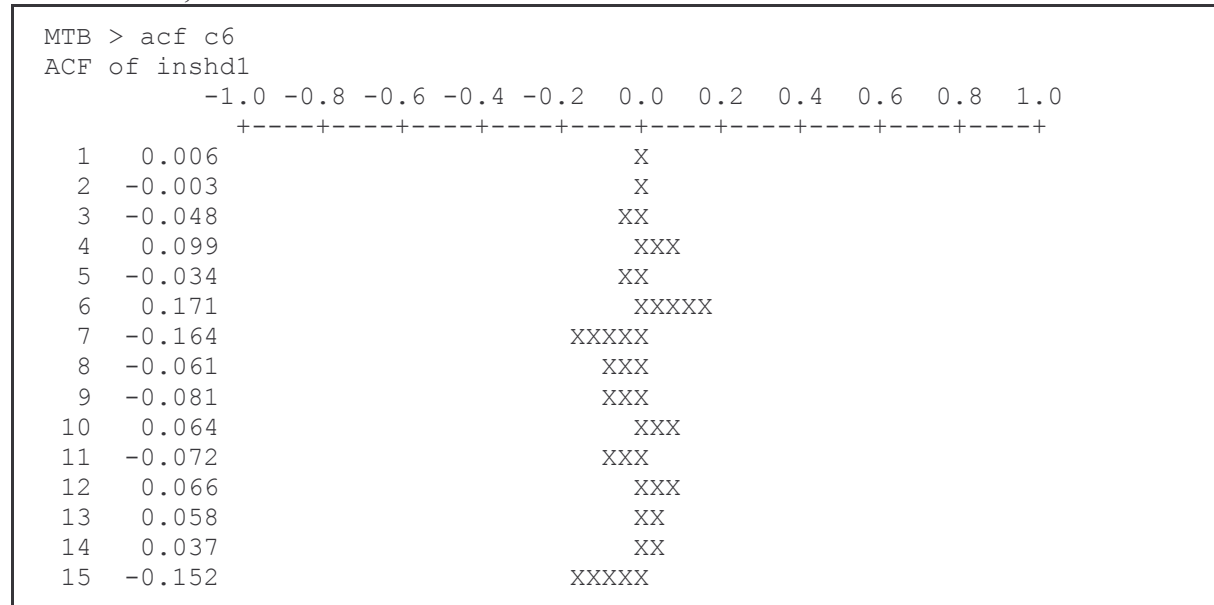
```
MTB > differences 1 c4 c6
MTB > name c6 'inshd1'
MTB > print c4 c6
```

ROW	inshore	inshd1
1	159492	*
2	157286	-2206
3	119363	-37923
4	138511	19148
5	144548	6037
6	131328	-13220
7	110527	-20801
8	110843	316
9	101859	-8984
10	101037	-822
11	97224	-3813
12	76588	-20636
13	62539	-14049
14	62052	-487
15	41648	-20404
16	35181	-6467
17	41213	6032
18	59939	18726
19	72623	12684
20	81455	8832
21	85822	4367
22	96523	10701
23	80038	-16485
24	113049	33011
25	106423	-6626
26	97721	-8702
27	79883	-17838
28	72369	-7514
29	78747	6378
30	101925	23178

To solve the problem take the differences from one year to the next, in the response variable (inshore catch). Taking the difference usually reduces the autocorrelation.



To check this, examine autocorrelation of the differenced variable.



Autocorrelation in response variable is usually reduced by taking differences.

Now examine whether **change** in the inshore catch (inshore catch after differencing) is related to offshore catch.

```
MTB > regress c6 1 c5;
SUBC> residuals c9.
```

The regression equation is  $\text{inshd1} = -4333 + 0.0603 \text{ canoff}$

29 cases used 1 cases contain missing values *(1956 lost from analysis)*

Predictor	Coef	Stdev	t-ratio	p
Constant	-4333	3798	-1.14	0.264
canoff	0.06033	0.06364	0.95	0.352

s = 15509      R-sq = 3.2%      R-sq(adj) = 0.0%

#### Analysis of Variance

SOURCE	DF	SS	MS	F	p
Regression	1	216159680	216159680	0.90	0.352
Error	27	6493937152	240516192		
Total	28	6710096896			

#### Unusual Observations

Obs.	canoff	inshd1	Fit	Stdev.Fit	Residual	St.Resid
3	4676	-37923	-4051	3611	-33872	-2.25R
24	94457	33011	1366	4559	31645	2.13R

Check the residuals for autocorrelation.

```
MTB > acf c9
```

ACF of C9

	-1.0	-0.8	-0.6	-0.4	-0.2	0.0	0.2	0.4	0.6	0.8	1.0
	+-----+-----+-----+-----+-----+-----+-----+-----+-----+-----+										
1	-0.002					X					
2	0.001					X					
3	-0.070					XXX					
4	0.051					XX					
5	-0.103					XXXX					
6	0.095					XXX					
7	-0.224					XXXXXXXX					
8	-0.130					XXXX					
9	-0.132					XXXX					
10	0.031					XX					
11	-0.090					XXX					
12	0.077					XXX					
13	0.095					XXX					
14	0.094					XXX					
15	-0.094					XXX					

Residuals no longer autocorrelated for new model (based on differencing)

**Conclusion:** When we remove the autocorrelation present in the inshore catch series, we find that the inshore catches are not related to offshore catches.

## Numerical Methods. Finding the sample size (srex9\_6.out)

Exercise 9.6 from Sokal and Rohlf (1995), page 268

What sample size should be used to be 80% certain of observing a true difference between two means as small as a tenth of a millimeter, at the 5% level of significance?

First compute the error Mean square = 0.2496

This is better estimate than total variance =  $25.6819/99 = 0.2594$

```
MTB > read 'srex9_5.dat' c1-c5;
SUBC> nob=20.
MTB > stack c1-c5 c6;
SUBC> subscripts c7.
MTB > name c6 'b_length' c7 'gr'
MTB > anova c6 = c7
```

Analysis of Variance for b\_length

Source	DF	SS	MS	F	P
gr	4	1.9734	0.4933	1.98	0.104
Error	95	23.7085	0.2496		
Total	99	25.6819			

$n = \text{unknown}$

$\sigma^2$  estimated as  $s^2 = 0.2496$  (see above)

$\delta = 0.10$  and  $\delta^2 = 0.01$

$v = a(n - 1)$

$\alpha = 5\%$

$P = 80\%$

match cdf computations in Minitab to t-values for example in Box 9.14 page 263

$t_{0.05[v]} = 2.642$  in text, for  $v = 4(20 - 1) = 76$

$t_{2(1-0.80)[v]} = 0.847$  in text, for  $v = 4(20 - 1) = 76$

```
MTB > invcdf .01;
SUBC> t 76.
0.0100 -2.3764
MTB > invcdf .005;
SUBC> t 76.
0.0050 -2.6421
MTB > invcdf .4;
SUBC> t 76.
0.4000 -0.2542
MTB > invcdf .2;
SUBC> t 76.
0.2000 -0.8464
```

use 0.005 and 0.20 for box 9.14

Use 0.005 and 0.20 for box 9.14 therefore use 0.025 and 0.20 for exercise 9.6

Compute  $k1 = 2(\sigma/\delta)^2$

```
MTB > let k1 = 2*(0.2496)/(0.01)
```

Guess  $n = 20$ , hence  $v = 2*(20-1) =$

38

```
MTB > invcdf 0.025 k2;
SUBC> t 38.
MTB > invcdf 0.2 k3;
SUBC> t 38.
MTB > let k4 = k1*(k2 + k3)**2      ≤ n
MTB > print k1 k2 k3 k4
K1      49.9200
K2      -2.02439
K3      -0.851178
K4      412.782      ≤ n
```

t value stored into k2

t value stored into k3

≤ n in Box 9.14

Both t-values are negative, the sum becomes positive when squared.

```
MTB > invcdf 0.025 k2;
SUBC> t 822.
MTB > invcdf 0.2 k3;
SUBC> t 822.
MTB > let k4 = k1*(k2 + k3)**2
MTB > print k2 k3 k4
K2      -1.96285
K3      -0.842055
K4      392.745      ≤ n
```

Guess  $n = 412$   
hence  $v = 822$

```
MTB > invcdf .025 k2;
SUBC> t 782.
MTB > invcdf .2 k3;
SUBC> t 782.
MTB > let k4 = k1*(k2 + k3)**2
MTB > print k4 k3 k2
K4      392.804      = n
K3      -0.842103
K2      -1.96301
MTB > stop
```

Guess  $n = 392$   
hence  $v = 782$

No change from last iteration

Sample size is  $n = 392$  for stated power and Type I error (= 5%).